

# Package ‘BAS’

April 26, 2022

**Version** 1.6.2

**Date** 2022-04-23

**Title** Bayesian Variable Selection and Model Averaging using Bayesian Adaptive Sampling

**Depends** R (>= 3.0)

**Imports** stats, graphics, utils, grDevices

**Suggests** MASS, knitr, ggplot2, GGally, rmarkdown, roxygen2, dplyr, glmbb, testthat, covr

**Description** Package for Bayesian Variable Selection and Model Averaging in linear models and generalized linear models using stochastic or deterministic sampling without replacement from posterior distributions. Prior distributions on coefficients are from Zellner's g-prior or mixtures of g-priors corresponding to the Zellner-Siow Cauchy Priors or the mixture of g-priors from Liang et al (2008) <[DOI:10.1198/016214507000001337](https://doi.org/10.1198/016214507000001337)> for linear models or mixtures of g-priors from Li and Clyde (2019) <[DOI:10.1080/01621459.2018.1469992](https://doi.org/10.1080/01621459.2018.1469992)> in generalized linear models. Other model selection criteria include AIC, BIC and Empirical Bayes estimates of g. Sampling probabilities may be updated based on the sampled models using sampling w/out replacement or an efficient MCMC algorithm which samples models using a tree structure of the model space as an efficient hash table. See Clyde, Ghosh and Littman (2010) <[DOI:10.1198/jcgs.2010.09049](https://doi.org/10.1198/jcgs.2010.09049)> for details on the sampling algorithms. Uniform priors over all models or beta-binomial prior distributions on model size are allowed, and for large p truncated priors on the model space may be used to enforce sampling models that are full rank. The user may force variables to always be included in addition to imposing constraints that higher order interactions are included only if their parents are included in the model. This material is based upon work supported by the National Science Foundation under Division of Mathematical Sciences grant 1106891. Any opinions, findings, and conclusions or recommendations expressed in this material are those of

the author(s) and do not necessarily reflect the views of the  
National Science Foundation.

**License** GPL (>= 3)

**URL** <https://www.r-project.org>, <https://github.com/merliseclyde/BAS>

**BugReports** <https://github.com/merliseclyde/BAS/issues>

**Repository** CRAN

**NeedsCompilation** yes

**ByteCompile** yes

**VignetteBuilder** knitr

**Encoding** UTF-8

**RoxygenNote** 7.1.2

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**Date/Publication** 2022-04-26 07:30:06 UTC

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**Index****81****Description**

Package for Bayesian Model Averaging in linear models using stochastic or deterministic sampling without replacement from posterior distributions. Prior distributions on coefficients are of the form of Zellner's g-prior or mixtures of g-priors. Options include the Zellner-Siow Cauchy Priors, the Liang et al hyper-g priors, Local and Global Empirical Bayes estimates of g, and other default

model selection criteria such as AIC and BIC. Sampling probabilities may be updated based on the sampled models.

### Details

Package: BAS  
Depends: R (>= 2.8)  
License: GPL-3  
URL: <https://www.stat.duke.edu/~clyde>

Index:

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### References

Clyde, M. Ghosh, J. and Littman, M. (2010) Bayesian Adaptive Sampling for Variable Selection and Model Averaging. *Journal of Computational Graphics and Statistics*. 20:80-101  
doi: [10.1198/jcgs.2010.09049](https://doi.org/10.1198/jcgs.2010.09049)

Clyde, M. and George, E. I. (2004) Model uncertainty. *Statist. Sci.*, 19, 81-94.  
doi: [10.1214/088342304000000035](https://doi.org/10.1214/088342304000000035)

Clyde, M. (1999) Bayesian Model Averaging and Model Search Strategies (with discussion). In *Bayesian Statistics 6*. J.M. Bernardo, A.P. Dawid, J.O. Berger, and A.F.M. Smith eds. Oxford University Press, pages 157-185.

Li, Y. and Clyde, M. (2018) Mixtures of g-priors in Generalized Linear Models. *Journal of the American Statistical Association*, 113:524, 1828-1845 doi: [10.1080/01621459.2018.1469992](https://doi.org/10.1080/01621459.2018.1469992)

Liang, F., Paulo, R., Molina, G., Clyde, M. and Berger, J.O. (2008) Mixtures of g-priors for Bayesian Variable Selection. *Journal of the American Statistical Association*. 103:410-423.  
doi: [10.1198/016214507000001337](https://doi.org/10.1198/016214507000001337)

### See Also

[bas.lm](#) [bas.glm](#)

Other bas methods: [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.hereditiy.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

**Examples**

```
data("Hald")
hald.gprior = bas.lm(Y ~ ., data=Hald, alpha=13, prior="g-prior")

# more complete demos

demo(BAS.hald)
## Not run:
demo(BAS.USCrime)

## End(Not run)
```

---

bas.glm

*Bayesian Adaptive Sampling Without Replacement for Variable Selection in Generalized Linear Models*

---

**Description**

Sample with or without replacement from a posterior distribution on GLMs

**Usage**

```
bas.glm(  
  formula,  
  family = binomial(link = "logit"),  
  data,  
  weights,  
  subset,  
  contrasts = NULL,  
  offset,  
  na.action = "na.omit",  
  n.models = NULL,  
  betaprior = CCH(alpha = 0.5, beta = as.numeric(nrow(data)), s = 0),  
  modelprior = beta.binomial(1, 1),  
  initprobs = "Uniform",  
  include.always = ~1,  
  method = "MCMC",  
  update = NULL,  
  bestmodel = NULL,  
  prob.rw = 0.5,  
  MCMC.iterations = NULL,  
  thin = 1,  
  control = glm.control(),  
  laplace = FALSE,  
  renormalize = FALSE,  
  force.heredity = FALSE,  
  bigmem = FALSE  
)
```

**Arguments**

formula	generalized linear model formula for the full model with all predictors, $Y \sim X$ . All code assumes that an intercept will be included in each model.
family	a description of the error distribution and link function for exponential family; currently only 'binomial()' with the logistic link and 'poisson()' with the log link are available.
data	data frame
weights	optional vector of weights to be used in the fitting process. May be missing in which case weights are 1.
subset	subset of data used in fitting
contrasts	an optional list. See the contrasts.arg of 'model.matrix.default()'.
offset	a priori known component to be included in the linear predictor; by default 0.
na.action	a function which indicates what should happen when the data contain NAs. The default is "na.omit".
n.models	number of unique models to keep. If NULL, BAS will attempt to enumerate unless $p > 35$ or method="MCMC". For any of methods using MCMC algorithms that sample with replacement, sampling will stop when the number of iterations exceeds the min of 'n.models' or 'MCMC.iterations' and on exit 'n.models' is updated to reflect the unique number of models that have been sampled.
betaprior	Prior on coefficients for model coefficients (except intercept). Options include <a href="#">g.prior</a> , <a href="#">CCH</a> , <a href="#">robust</a> , <a href="#">intrinsic</a> , <a href="#">beta.prime</a> , <a href="#">EB.local</a> , <a href="#">AIC</a> , and <a href="#">BIC</a> .
modelprior	Family of prior distribution on the models. Choices include <a href="#">uniform</a> , <a href="#">Bernoulli</a> , <a href="#">beta.binomial</a> , truncated Beta-Binomial, <a href="#">tr.beta.binomial</a> , and truncated power family <a href="#">tr.power.prior</a> .
initprobs	vector of length $p$ with the initial inclusion probabilities used for sampling without replacement (the intercept will be included with probability one and does not need to be added here) or a character string giving the method used to construct the sampling probabilities if "Uniform" each predictor variable is equally likely to be sampled (equivalent to random sampling without replacement). If "eplogp", use the <a href="#">eplogprob</a> function to approximate the Bayes factor using p-values to find initial marginal inclusion probabilities and sample without replacement using these inclusion probabilities, which may be updated using estimates of the marginal inclusion probabilities. "eplogp" assumes that MLEs from the full model exist; for problems where that is not the case or 'p' is large, initial sampling probabilities may be obtained using <a href="#">eplogprob.marg</a> which fits a model to each predictor separately. To run a Markov Chain to provide initial estimates of marginal inclusion probabilities, use method="MCMC+BAS" below. While the initprobs are not used in sampling for method="MCMC", this determines the order of the variables in the lookup table and affects memory allocation in large problems where enumeration is not feasible. For variables that should always be included set the corresponding initprobs to 1, to override the 'modelprior' or use 'include.always' to force these variables to always be included in the model.

include.always	A formula with terms that should always be included in the model with probability one. By default this is '~ 1' meaning that the intercept is always included. This will also override any of the values in 'initprobs' above by setting them to 1.
method	A character variable indicating which sampling method to use: method="BAS" uses Bayesian Adaptive Sampling (without replacement) using the sampling probabilities given in initprobs and updates using the marginal inclusion probabilities to direct the search/sample; method="MCMC" combines a random walk Metropolis Hastings (as in MC3 of Raftery et al 1997) with a random swap of a variable included with a variable that is currently excluded (see Clyde, Ghosh, and Littman (2010) for details); method="MCMC+BAS" runs an initial MCMC as above to calculate marginal inclusion probabilities and then samples without replacement as in BAS; method = "deterministic" runs a deterministic sampling using the initial probabilities (no updating); this is recommended for fast enumeration or if a model of independence is a good approximation to the joint posterior distribution of the model indicators. For BAS, the sampling probabilities can be updated as more models are sampled. (see 'update' below). We recommend "MCMC+BAS" or "MCMC" for high dimensional problems.
update	number of iterations between potential updates of the sampling probabilities in the "BAS" method. If NULL do not update, otherwise the algorithm will update using the marginal inclusion probabilities as they change while sampling takes place. For large model spaces, updating is recommended. If the model space will be enumerated, leave at the default.
bestmodel	optional binary vector representing a model to initialize the sampling. If NULL sampling starts with the null model
prob.rw	For any of the MCMC methods, probability of using the random-walk proposal; otherwise use a random "flip" move to propose a new model.
MCMC.iterations	Number of models to sample when using any of the MCMC options; should be greater than 'n.models'. By default 10*n.models.
thin	oFr "MCMC", thin the MCMC chain every "thin" iterations; default is no thinning. For large p, thinning can be used to significantly reduce memory requirements as models and associated summaries are saved only every thin iterations. For thin = p, the model and associated output are recorded every p iterations, similar to the Gibbs sampler in SSVS.
control	a list of parameters that control convergence in the fitting process. See the documentation for glm.control()
laplace	logical variable for whether to use a Laplace approximate for integration with respect to g to obtain the marginal likelihood. If FALSE the Cephes library is used which may be inaccurate for large n or large values of the Wald Chisquared statistic.
renormalize	logical variable for whether posterior probabilities should be based on renormalizing marginal likelihoods times prior probabilities or use Monte Carlo frequencies. Applies only to MCMC sampling.
force.heredity	Logical variable to force all levels of a factor to be included together and to include higher order interactions only if lower order terms are included. Currently

	only supported with 'method='MCMC'' and 'method='BAS'' (experimental) on non-Solaris platforms. Default is FALSE.
bigmem	Logical variable to indicate that there is access to large amounts of memory (physical or virtual) for enumeration with large model spaces, e.g. $> 2^{25}$ .

### Details

BAS provides several search algorithms to find high probability models for use in Bayesian Model Averaging or Bayesian model selection. For  $p$  less than 20-25, BAS can enumerate all models depending on memory availability, for larger  $p$ , BAS samples without replacement using random or deterministic sampling. The Bayesian Adaptive Sampling algorithm of Clyde, Ghosh, Littman (2010) samples models without replacement using the initial sampling probabilities, and will optionally update the sampling probabilities every "update" models using the estimated marginal inclusion probabilities. BAS uses different methods to obtain the `initprobs`, which may impact the results in high-dimensional problems. The deterministic sampler provides a list of the top models in order of an approximation of independence using the provided `initprobs`. This may be effective after running the other algorithms to identify high probability models and works well if the correlations of variables are small to modest. The priors on coefficients are mixtures of g-priors that provide approximations to the power prior.

### Value

`bas.glm` returns an object of class `basglm`

An object of class `basglm` is a list containing at least the following components:

<code>postprobs</code>	the posterior probabilities of the models selected
<code>priorprobs</code>	the prior probabilities of the models selected
<code>logmarg</code>	values of the log of the marginal likelihood for the models
<code>n.vars</code>	total number of independent variables in the full model, including the intercept
<code>size</code>	the number of independent variables in each of the models, includes the intercept
<code>which</code>	a list of lists with one list per model with variables that are included in the model
<code>probne0</code>	the posterior probability that each variable is non-zero
<code>coefficients</code>	list of lists with one list per model giving the GLM estimate of each (nonzero) coefficient for each model.
<code>se</code>	list of lists with one list per model giving the GLM standard error of each coefficient for each model
<code>deviance</code>	the GLM deviance for each model
<code>modelprior</code>	the prior distribution on models that created the BMA object
<code>Q</code>	the Q statistic for each model used in the marginal likelihood approximation
<code>Y</code>	response
<code>X</code>	matrix of predictors
<code>family</code>	family object from the original call
<code>betaprior</code>	family object for prior on coefficients, including hyperparameters
<code>modelprior</code>	family object for prior on the models
<code>include.always</code>	indices of variables that are forced into the model



**Author(s)**

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**References**

- Li, Y. and Clyde, M. (2018) Mixtures of g-priors in Generalized Linear Models. *Journal of the American Statistical Association*. 113:1828-1845  
doi: [10.1080/01621459.2018.1469992](https://doi.org/10.1080/01621459.2018.1469992)
- Clyde, M. Ghosh, J. and Littman, M. (2010) Bayesian Adaptive Sampling for Variable Selection and Model Averaging. *Journal of Computational Graphics and Statistics*. 20:80-101  
doi: [10.1198/jcgs.2010.09049](https://doi.org/10.1198/jcgs.2010.09049)
- Raftery, A.E, Madigan, D. and Hoeting, J.A. (1997) Bayesian Model Averaging for Linear Regression Models. *Journal of the American Statistical Association*.

**Examples**

```
library(MASS)
data(Pima.tr)

# enumeration with default method="BAS"
pima.cch = bas.glm(type ~ ., data=Pima.tr, n.models= 2^7,
                  method="BAS",
                  betaprior=CCH(a=1, b=532/2, s=0), family=binomial(),
                  modelprior=beta.binomial(1,1))

summary(pima.cch)
image(pima.cch)

# Note MCMC.iterations are set to 1000 for illustration purposes
# Please check convergence diagnostics and run longer in practice

pima.robust = bas.glm(type ~ ., data=Pima.tr, n.models= 2^7,
                    method="MCMC", MCMC.iterations=5000,
                    betaprior=robust(), family=binomial(),
                    modelprior=beta.binomial(1,1))

pima.BIC = bas.glm(type ~ ., data=Pima.tr, n.models= 2^7,
                  method="BAS+MCMC", MCMC.iterations=5000,
                  betaprior=bic.prior(), family=binomial(),
                  modelprior=uniform())

# Poisson example
if(requireNamespace("glmbb", quietly=TRUE)) {
  data(crabs, package='glmbb')
  #short run for illustration
  crabs.bas = bas.glm(satell ~ color*spine*width + weight, data=crabs,
                    family=poisson(),
                    betaprior=EB.local(), modelprior=uniform(),
                    method='MCMC', n.models=2^10, MCMC.iterations=5000,
                    prob.rw=.95)
```

```
}
```

---

bas.lm

*Bayesian Adaptive Sampling for Bayesian Model Averaging and Variable Selection in Linear Models*

---

## Description

Sample without replacement from a posterior distribution on models

## Usage

```
bas.lm(  
  formula,  
  data,  
  subset,  
  weights,  
  contrasts = NULL,  
  na.action = "na.omit",  
  n.models = NULL,  
  prior = "ZS-null",  
  alpha = NULL,  
  modelprior = beta.binomial(1, 1),  
  initprobs = "Uniform",  
  include.always = ~1,  
  method = "BAS",  
  update = NULL,  
  bestmodel = NULL,  
  prob.local = 0,  
  prob.rw = 0.5,  
  MCMC.iterations = NULL,  
  lambda = NULL,  
  delta = 0.025,  
  thin = 1,  
  renormalize = FALSE,  
  force.heredity = FALSE,  
  pivot = TRUE,  
  tol = 1e-07,  
  bigmem = FALSE  
)
```

## Arguments

formula      linear model formula for the full model with all predictors,  $Y \sim X$ . All code assumes that an intercept will be included in each model and that the X's will be centered.

data	a data frame. Factors will be converted to numerical vectors based on the using 'model.matrix'.
subset	an optional vector specifying a subset of observations to be used in the fitting process.
weights	an optional vector of weights to be used in the fitting process. Should be NULL or a numeric vector. If non-NULL, Bayes estimates are obtained assuming that $Y \sim N(Xb, \sigma^2 \text{diag}(1/\text{weights}))$ .
contrasts	an optional list. See the contrasts.arg of 'model.matrix.default()'.
na.action	a function which indicates what should happen when the data contain NAs. The default is "na.omit".
n.models	number of models to sample either without replacement (method="BAS" or "MCMC+BAS") or with replacement (method="MCMC"). If NULL, BAS with method="BAS" will try to enumerate all $2^p$ models. If enumeration is not possible (memory or time) then a value should be supplied which controls the number of sampled models using 'n.models'. With method="MCMC", sampling will stop once the $\min(\text{n.models}, \text{MCMC.iterations})$ occurs so MCMC.iterations be significantly larger than n.models in order to explore the model space. On exit for method= "MCMC" this is the number of unique models that have been sampled with counts stored in the output as "freq".
prior	prior distribution for regression coefficients. Choices include <ul style="list-style-type: none"> <li>• "AIC"</li> <li>• "BIC"</li> <li>• "g-prior", Zellner's g prior where 'g' is specified using the argument 'alpha'</li> <li>• "JZS" Jeffreys-Zellner-Siow prior which uses the Jeffreys prior on sigma and the Zellner-Siow Cauchy prior on the coefficients. The optional parameter 'alpha' can be used to control the squared scale of the prior, where the default is alpha=1. Setting 'alpha' is equal to <math>\text{rscale}^2</math> in the BayesFactor package of Morey. This uses QUADMATH for numerical integration of g.</li> <li>• "ZS-null", a Laplace approximation to the 'JZS' prior for integration of g. alpha = 1 only. We recommend using 'JZS' for accuracy and compatibility with the BayesFactor package, although it is slower.</li> <li>• "ZS-full" (to be deprecated)</li> <li>• "hyper-g", a mixture of g-priors where the prior on <math>g/(1+g)</math> is a <math>\text{Beta}(1, \text{alpha}/2)</math> as in Liang et al (2008). This uses the Cephes library for evaluation of the marginal likelihoods and may be numerically unstable for large n or R2 close to 1. Default choice of alpha is 3.</li> <li>• "hyper-g-laplace", Same as above but using a Laplace approximation to integrate over the prior on g.</li> <li>• "hyper-g-n", a mixture of g-priors that where <math>u = g/n</math> and <math>u \sim \text{Beta}(1, \text{alpha}/2)</math> to provide consistency when the null model is true.</li> <li>• "EB-local", use the MLE of g from the marginal likelihood within each model</li> <li>• "EB-global" uses an EM algorithm to find a common or global estimate of g, averaged over all models. When it is not possible to enumerate all models, the EM algorithm uses only the models sampled under EB-local.</li> </ul>

alpha	optional hyperparameter in g-prior or hyper g-prior. For Zellner's g-prior, alpha = g, for the Liang et al hyper-g or hyper-g-n method, recommended choice is alpha are between ( $2 < \alpha < 4$ ), with alpha = 3 the default. For the Zellner-Siow prior alpha = 1 by default, but can be used to modify the rate parameter in the gamma prior on g, $1/g \sim G(1/2, n*\alpha/2)$ so that $\beta \sim C(0, \sigma^2 \alpha (X'X/n)^{-1})$ .
modelprior	Family of prior distribution on the models. Choices include <code>uniformBernoulli</code> or <code>beta.binomial</code> , <code>tr.beta.binomial</code> , (with truncation) <code>tr.poisson</code> (a truncated Poisson), and <code>tr.power.prior</code> (a truncated power family), with the default being a <code>beta.binomial(1, 1)</code> . Truncated versions are useful for $p > n$ .
initprobs	Vector of length p or a character string specifying which method is used to create the vector. This is used to order variables for sampling all methods for potentially more efficient storage while sampling and provides the initial inclusion probabilities used for sampling without replacement with <code>method="BAS"</code> . Options for the character string giving the method are: "Uniform" or "uniform" where each predictor variable is equally likely to be sampled (equivalent to random sampling without replacement); "eplogp" uses the <code>eplogprob</code> function to approximate the Bayes factor from p-values from the full model to find initial marginal inclusion probabilities; "marg-eplogp" uses <code>eplogprob.marg</code> function to approximate the Bayes factor from p-values from the full model each simple linear regression. To run a Markov Chain to provide initial estimates of marginal inclusion probabilities for "BAS", use <code>method="MCMC+BAS"</code> below. While the <code>initprobs</code> are not used in sampling for <code>method="MCMC"</code> , this determines the order of the variables in the lookup table and affects memory allocation in large problems where enumeration is not feasible. For variables that should always be included set the corresponding <code>initprobs</code> to 1, to override the 'modelprior' or use 'include.always' to force these variables to always be included in the model.
include.always	A formula with terms that should always be included in the model with probability one. By default this is '~ 1' meaning that the intercept is always included. This will also override any of the values in 'initprobs' above by setting them to 1.
method	A character variable indicating which sampling method to use: <ul style="list-style-type: none"> <li>"deterministic" uses the "top k" algorithm described in Ghosh and Clyde (2011) to sample models in order of approximate probability under conditional independence using the "initprobs". This is the most efficient algorithm for enumeration.</li> <li>"BAS" uses Bayesian Adaptive Sampling (without replacement) using the sampling probabilities given in <code>initprobs</code> under a model of conditional independence. These can be updated based on estimates of the marginal inclusion probabilities.</li> <li>"MCMC" samples with replacement via a MCMC algorithm that combines the birth/death random walk in Hoeting et al (1997) of MC3 with a random swap move to interchange a variable in the model with one currently excluded as described in Clyde, Ghosh and Littman (2010).</li> <li>"MCMC+BAS" runs an initial MCMC to calculate marginal inclusion probabilities and then samples without replacement as in BAS. For BAS, the</li> </ul>

sampling probabilities can be updated as more models are sampled. (see update below).

update	number of iterations between potential updates of the sampling probabilities for method "BAS" or "MCMC+BAS". If NULL do not update, otherwise the algorithm will update using the marginal inclusion probabilities as they change while sampling takes place. For large model spaces, updating is recommended. If the model space will be enumerated, leave at the default.
bestmodel	optional binary vector representing a model to initialize the sampling. If NULL sampling starts with the null model
prob.local	A future option to allow sampling of models "near" the median probability model. Not used at this time.
prob.rw	For any of the MCMC methods, probability of using the random-walk Metropolis proposal; otherwise use a random "flip" move to propose swap a variable that is excluded with a variable in the model.
MCMC.iterations	Number of iterations for the MCMC sampler; the default is $n.models * 10$ if not set by the user.
lambda	Parameter in the AMCMC algorithm (deprecated).
delta	truncation parameter to prevent sampling probabilities to degenerate to 0 or 1 prior to enumeration for sampling without replacement.
thin	For "MCMC", thin the MCMC chain every "thin" iterations; default is no thinning. For large p, thinning can be used to significantly reduce memory requirements as models and associated summaries are saved only every thin iterations. For thin = p, the model and associated output are recorded every p iterations, similar to the Gibbs sampler in SSVS.
renormalize	For MCMC sampling, should posterior probabilities be based on renormalizing the marginal likelihoods times prior probabilities (TRUE) or frequencies from MCMC. The latter are unbiased in long runs, while the former may have less variability. May be compared via the diagnostic plot function <a href="#">diagnostics</a> . See details in Clyde and Ghosh (2012).
force.heredity	Logical variable to force all levels of a factor to be included together and to include higher order interactions only if lower order terms are included. Currently supported with 'method='MCMC'' and experimentally with 'method='BAS'' on non-Solaris platforms. Default is FALSE.
pivot	Logical variable to allow pivoting of columns when obtaining the OLS estimates of a model so that models that are not full rank can be fit. Defaults to TRUE. Currently coefficients that are not estimable are set to zero. Use caution with interpreting BMA estimates of parameters.
tol	1e-7 as
bigmem	Logical variable to indicate that there is access to large amounts of memory (physical or virtual) for enumeration with large model spaces, e.g. $> 2^{25}$ . default; used in determining rank of $X^T X$ in cholesky decomposition with pivoting.

## Details

BAS provides several algorithms to sample from posterior distributions of models for use in Bayesian Model Averaging or Bayesian variable selection. For  $p$  less than 20-25, BAS can enumerate all models depending on memory availability. As BAS saves all models, MLEs, standard errors, log marginal likelihoods, prior and posterior and probabilities memory requirements grow linearly with  $M \cdot p$  where  $M$  is the number of models and  $p$  is the number of predictors. For example, enumeration with  $p=21$  with 2,097,152 takes just under 2 Gigabytes on a 64 bit machine to store all summaries that would be needed for model averaging. (A future version will likely include an option to not store all summaries if users do not plan on using model averaging or model selection on Best Predictive models.) For larger  $p$ , BAS samples without replacement using random or deterministic sampling. The Bayesian Adaptive Sampling algorithm of Clyde, Ghosh, Littman (2010) samples models without replacement using the initial sampling probabilities, and will optionally update the sampling probabilities every "update" models using the estimated marginal inclusion probabilities. BAS uses different methods to obtain the `initprobs`, which may impact the results in high-dimensional problems. The deterministic sampler provides a list of the top models in order of an approximation of independence using the provided `initprobs`. This may be effective after running the other algorithms to identify high probability models and works well if the correlations of variables are small to modest. We recommend "MCMC" for problems where enumeration is not feasible (memory or time constrained) or even modest  $p$  if the number of models sampled is not close to the number of possible models and/or there are significant correlations among the predictors as the bias in estimates of inclusion probabilities from "BAS" or "MCMC+BAS" may be large relative to the reduced variability from using the normalized model probabilities as shown in Clyde and Ghosh, 2012. Diagnostic plots with MCMC can be used to assess convergence. For large problems we recommend thinning with MCMC to reduce memory requirements. The priors on coefficients include Zellner's  $g$ -prior, the Hyper- $g$  prior (Liang et al 2008, the Zellner-Siow Cauchy prior, Empirical Bayes (local and global)  $g$ -priors. AIC and BIC are also included, while a range of priors on the model space are available.

## Value

`bas` returns an object of class `bas`

An object of class `BAS` is a list containing at least the following components:

<code>postprob</code>	the posterior probabilities of the models selected
<code>priorprobs</code>	the prior probabilities of the models selected
<code>namesx</code>	the names of the variables
<code>R2</code>	R2 values for the models
<code>logmarg</code>	values of the log of the marginal likelihood for the models. This is equivalent to the log Bayes Factor for comparing each model to a base model with intercept only.
<code>n.vars</code>	total number of independent variables in the full model, including the intercept
<code>size</code>	the number of independent variables in each of the models, includes the intercept
<code>rank</code>	the rank of the design matrix; if <code>'pivot = FALSE'</code> , this is the same as <code>size</code> as no checking of rank is conducted.
<code>which</code>	a list of lists with one list per model with variables that are included in the model

<code>probne0</code>	the posterior probability that each variable is non-zero computed using the renormalized marginal likelihoods of sampled models. This may be biased if the number of sampled models is much smaller than the total number of models. Unbiased estimates may be obtained using method "MCMC".
<code>mle</code>	list of lists with one list per model giving the MLE (OLS) estimate of each (nonzero) coefficient for each model. NOTE: The intercept is the mean of Y as each column of X has been centered by subtracting its mean.
<code>mle.se</code>	list of lists with one list per model giving the MLE (OLS) standard error of each coefficient for each model
<code>prior</code>	the name of the prior that created the BMA object
<code>alpha</code>	value of hyperparameter in coefficient prior used to create the BMA object.
<code>modelprior</code>	the prior distribution on models that created the BMA object
<code>Y</code>	response
<code>X</code>	matrix of predictors
<code>mean.x</code>	vector of means for each column of X (used in <code>predict.bas</code> )
<code>include.always</code>	indices of variables that are forced into the model

The function `summary.bas`, is used to print a summary of the results. The function `plot.bas` is used to plot posterior distributions for the coefficients and `image.bas` provides an image of the distribution over models. Posterior summaries of coefficients can be extracted using `coefficients.bas`. Fitted values and predictions can be obtained using the S3 functions `fitted.bas` and `predict.bas`. BAS objects may be updated to use a different prior (without rerunning the sampler) using the function `update.bas`. For MCMC sampling `diagnostics` can be used to assess whether the MCMC has run long enough so that the posterior probabilities are stable. For more details see the associated demos and vignette.

### Author(s)

Merlise Clyde (<clyde@duke.edu>) and Michael Littman

### References

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Rouder, J. N., Morey, R. D., Speckman, P. L., Province, J. M., (2012) Default Bayes Factors for ANOVA Designs. *Journal of Mathematical Psychology*. 56. p. 356-374.

### See Also

[summary.bas](#), [coefficients.bas](#), [print.bas](#), [predict.bas](#), [fitted.bas](#), [plot.bas](#), [image.bas](#), [eplogprob](#), [update.bas](#)

Other bas methods: [BAS](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

### Examples

```
library(MASS)
data(UScrime)
crime.bic <- bas.lm(log(y) ~ log(M) + So + log(Ed) +
  log(Po1) + log(Po2) +
  log(LF) + log(M.F) + log(Pop) + log(NW) +
  log(U1) + log(U2) + log(GDP) + log(Ineq) +
  log(Prob) + log(Time),
data = UScrime, n.models = 2^15, prior = "BIC",
modelprior = beta.binomial(1, 1),
initprobs = "eplogp", pivot = FALSE
)

# use MCMC rather than enumeration
crime.mcmc <- bas.lm(log(y) ~ log(M) + So + log(Ed) +
  log(Po1) + log(Po2) +
  log(LF) + log(M.F) + log(Pop) + log(NW) +
  log(U1) + log(U2) + log(GDP) + log(Ineq) +
  log(Prob) + log(Time),
data = UScrime,
method = "MCMC",
MCMC.iterations = 20000, prior = "BIC",
modelprior = beta.binomial(1, 1),
initprobs = "eplogp", pivot = FALSE
)
```



```

summary(crime.bic)
plot(crime.bic)
image(crime.bic, subset = -1)

# example with two-way interactions and hierarchical constraints
data(ToothGrowth)
ToothGrowth$dose <- factor(ToothGrowth$dose)
levels(ToothGrowth$dose) <- c("Low", "Medium", "High")
TG.bas <- bas.lm(len ~ supp * dose,
  data = ToothGrowth,
  modelprior = uniform(), method = "BAS",
  force.heredity = TRUE
)
summary(TG.bas)
image(TG.bas)

# exmple with non-full rank case

loc <- system.file("testdata", package = "BAS")
d <- read.csv(paste(loc, "JASP-testdata.csv", sep = "/"))
fullModelFormula <- as.formula("contNormal ~ contGamma * contExpon +
  contGamma * contcor1 + contExpon * contcor1")

# should give warning; use pivot = T to fit non-full rank case (fails on i386)
out = bas.lm(fullModelFormula,
  data = d,
  alpha = 0.125316,
  prior = "JZS",
  weights = facFifty, force.heredity = FALSE, pivot = FALSE)

# use pivot = TRUE to fit non-full rank case (default)
# This is slower but safer

out = bas.lm(fullModelFormula,
  data = d,
  alpha = 0.125316,
  prior = "JZS",
  weights = facFifty, force.heredity = FALSE, pivot = TRUE)

# more complete demo's
demo(BAS.hald)
## Not run:
demo(BAS.USCrime)

## End(Not run)

```

**Description**

Calculate the posterior probability that the absolute value of error exceeds more than  $k$  standard deviations  $P(|\epsilon_j| > k \sigma \mid \text{data})$  under the model  $Y = X B + \epsilon$ , with  $\epsilon \sim N(0, \sigma^2 I)$  based on the paper by Chaloner & Brant *Biometrika* (1988). Either  $k$  or the prior probability of there being no outliers must be provided. This only uses the reference prior  $p(B, \sigma) = 1$ ; other priors and model averaging to come.

**Usage**

```
Bayes.outlier(lmobj, k, prior.prob)
```

**Arguments**

<code>lmobj</code>	An object of class 'lm'
<code>k</code>	number of standard deviations used in calculating probability of an individual case being an outlier, $P( \text{error}  > k \sigma \mid \text{data})$
<code>prior.prob</code>	The prior probability of there being no outliers in the sample of size $n$

**Value**

Returns a list of three items:

<code>e</code>	residuals
<code>hat</code>	leverage values
<code>prob.outlier</code>	posterior probabilities of a point being an outlier
<code>prior.prob</code>	prior probability of a point being an outlier

**References**

Chaloner & Brant (1988) A Bayesian Approach to Outlier Detection and Residual Analysis *Biometrika* (1988) 75, 651-659

**Examples**

```
data("stackloss")
stack.lm <- lm(stack.loss ~ ., data = stackloss)
stack.outliers <- Bayes.outlier(stack.lm, k = 3)
plot(stack.outliers$prob.outlier, type = "h", ylab = "Posterior Probability")
# adjust for sample size for calculating prior prob that a
# a case is an outlier
stack.outliers <- Bayes.outlier(stack.lm, prior.prob = 0.95)
# cases where posterior probability exceeds prior probability
which(stack.outliers$prob.outlier > stack.outliers$prior.prob)
```

---

bayesglm.fit	<i>Fitting Generalized Linear Models and Bayesian marginal likelihood evaluation</i>
--------------	--

---

### Description

A version of glm.fit rewritten in C; also returns marginal likelihoods for Bayesian model comparison

### Usage

```
bayesglm.fit(
  x,
  y,
  weights = rep(1, nobs),
  start = NULL,
  etastart = NULL,
  mustart = NULL,
  offset = rep(0, nobs),
  family = binomial(),
  coefprior = bic.prior(nobs),
  control = glm.control(),
  intercept = TRUE
)
```

### Arguments

x	design matrix
y	response
weights	optional vector of weights to be used in the fitting process. Should be NULL or a numeric vector.
start	starting value for coefficients in the linear predictor
etastart	starting values for the linear predictor
mustart	starting values for the vectors of means
offset	a priori known component to be included in the linear predictor
family	a description of the error distribution and link function for exponential family; currently only binomial() and poisson() with canonical links are implemented.
coefprior	function specifying prior distribution on coefficients with optional hyperparameters leading to marginal likelihood calculations; options include bic.prior(), aic.prior(), and ic.prior()
control	a list of parameters that control convergence in the fitting process. See the documentation for glm.control()
intercept	should an intercept be included in the null model?

**Details**

C version of glm-fit. For different prior choices returns, marginal likelihood of model using a Laplace approximation.

**Value**

coefficients	MLEs
se	Standard errors of coefficients based on the sqrt of the diagonal of the inverse information matrix
mu	fitted mean
rank	numeric rank of the fitted linear model
deviance	minus twice the log likelihood evaluated at the MLEs
g	value of g in g-priors
shrinkage	shrinkage factor for coefficients in linear predictor
RegSS	quadratic form $\beta'I(\beta)\beta$ used in shrinkage
logmarglik	the log marginal or integrated log likelihood (up to a constant)

**Author(s)**

Merlise Clyde translated the `glm.fit` from R base into C using the `.Call` interface

**References**

[glm](#)

**See Also**

[bic.prior](#)

**Examples**

```
data(Pima.tr, package="MASS")
Y <- as.numeric(Pima.tr$type) - 1
X <- cbind(1, as.matrix(Pima.tr[,1:7]))
out <- bayesglm.fit(X, Y, family=binomial(),coefprior=bic.prior(n=length(Y)))
out$coef
out$se
# using built in function
glm(type ~ ., family=binomial(), data=Pima.tr)
```

**Description**

Creates an object representing the prior distribution on models for BAS.

**Usage**

```
Bernoulli(probs = 0.5)
```

**Arguments**

probs	a scalar or vector of prior inclusion probabilities. If a scalar, the values is replicated for all variables and a 1 is added for the intercept. BAS checks to see if the length is equal to the dimension of the parameter vector for the full model and adds a 1 to include the intercept.
-------	--

**Details**

The independent Bernoulli prior distribution is a commonly used prior in BMA, with the Uniform distribution a special case with  $\text{probs}=.5$ . If all indicator variables have a independent Bernoulli distributions with common probability  $\text{probs}$ , the distribution on model size binomial( $p$ ,  $\text{probs}$ ) distribution.

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[bas.lm](#), [beta.binomial](#), [uniform](#)

Other priors modelpriors: [Bernoulli.heredity\(\)](#), [beta.binomial\(\)](#), [tr.beta.binomial\(\)](#), [tr.poisson\(\)](#), [tr.power.prior\(\)](#), [uniform\(\)](#)

**Examples**

```
Bernoulli(.9)
```

---

Bernoulli.heredity	<i>Independent Bernoulli prior on models that with constraints for model hierarchy induced by interactions</i>
--------------------	--

---

**Description**

Independent Bernoulli prior on models that with constraints for model hierarchy induced by interactions

**Usage**

```
Bernoulli.heredity(pi = 0.5, parents)
```

**Arguments**

pi	Bernoulli probability that term is included
parents	matrix of terms and parents with indicators of which terms are parents for each term

**Note**

Not implemented yet for use with bas.lm or bas.glm

**See Also**

Other priors modelpriors: [Bernoulli\(\)](#), [beta.binomial\(\)](#), [tr.beta.binomial\(\)](#), [tr.poisson\(\)](#), [tr.power.prior\(\)](#), [uniform\(\)](#)

---

beta.binomial	<i>Beta-Binomial Prior Distribution for Models</i>
---------------	--

---

**Description**

Creates an object representing the prior distribution on models for BAS.

**Usage**

```
beta.binomial(alpha = 1, beta = 1)
```

**Arguments**

alpha	parameter in the beta prior distribution
beta	parameter in the beta prior distribution

**Details**

The beta-binomial distribution on model size is obtained by assigning each variable inclusion indicator independent Bernoulli distributions with probability  $w$ , and then giving  $w$  a  $\text{beta}(\alpha, \beta)$  distribution. Marginalizing over  $w$  leads to the distribution on model size having the beta-binomial distribution. The default hyperparameters lead to a uniform distribution over model size.

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[bas.lm](#), [Bernoulli](#), [uniform](#)

Other priors modelpriors: [Bernoulli.heredity\(\)](#), [Bernoulli\(\)](#), [tr.beta.binomial\(\)](#), [tr.poisson\(\)](#), [tr.power.prior\(\)](#), [uniform\(\)](#)

**Examples**

```
beta.binomial(1, 10) #' @family priors modelpriors
```

---

beta.prior

*Beta-Prime Prior Distribution for Coefficients in BMA Model*

---

**Description**

Creates an object representing the Beta-Prime prior that is mixture of g-priors on coefficients for BAS.

**Usage**

```
beta.prior(n = NULL)
```

**Arguments**

`n` the sample size; if NULL, the value derived from the data in the call to 'bas.glm' will be used.

**Details**

Creates a structure used for [bas.glm](#).

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**[CCH](#)Other beta priors: [CCH\(\)](#), [EB.local\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [g.prior\(\)](#), [hyper.g.n\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [tCCH\(\)](#), [testBF.prior\(\)](#)**Examples**

```
beta.prime(n = 100)
```

---

`bodyfat`*Bodyfat Data*

---

**Description**

Lists estimates of the percentage of body fat determined by underwater weighing and various body circumference measurements for 252 men. Accurate measurement of body fat is inconvenient/costly and it is desirable to have easy methods of estimating body fat that are not inconvenient/costly.

**Format**

A data frame with 252 observations on the following 15 variables.

**Density** a numeric vector for the density determined from underwater weighing

**Bodyfat** percent body fat from Siri's (1956) equation

**Age** age of individual in years

**Weight** weight of the individual in pounds

**Height** height of individual in inches

**Neck** neck circumference in centimeters (cm)

**Chest** chest circumference (cm)

**Abdomen** abdomen circumference (cm)

**Hip** hip circumference (cm)

**"Thigh"** thigh circumference (cm)

**"Knee"** knee circumference (cm)

**Ankle** ankle circumference (cm)

**Biceps** bicep (extended) circumference (cm)

**Forearm** forearm circumference (cm)

**Wrist** wrist circumference (cm)



## Details

A variety of popular health books suggest that the readers assess their health, at least in part, by estimating their percentage of body fat. In Bailey (1994), for instance, the reader can estimate body fat from tables using their age and various skin-fold measurements obtained by using a caliper. Other texts give predictive equations for body fat using body circumference measurements (e.g. abdominal circumference) and/or skin-fold measurements. See, for instance, Behnke and Wilmore (1974), pp. 66-67; Wilmore (1976), p. 247; or Katch and McArdle (1977), pp. 120-132).#

Percentage of body fat for an individual can be estimated once body density has been determined. Folks (e.g. Siri (1956)) assume that the body consists of two components - lean body tissue and fat tissue. Letting

$D$  = Body Density ( $\text{gm}/\text{cm}^3$ )  $A$  = proportion of lean body tissue  $B$  = proportion of fat tissue ( $A+B=1$ )  $a$  = density of lean body tissue ( $\text{gm}/\text{cm}^3$ )  $b$  = density of fat tissue ( $\text{gm}/\text{cm}^3$ )

we have  $D = 1/[(A/a) + (B/b)]$  and solving for  $B$  we find  $B = (1/D)*[ab/(a-b)] - [b/(a-b)]$ .

Using the estimates  $a=1.10 \text{ gm}/\text{cm}^3$  and  $b=0.90 \text{ gm}/\text{cm}^3$  (see Katch and McArdle (1977), p. 111 or Wilmore (1976), p. 123) we come up with "Siri's equation":

Percentage of Body Fat (i.e.  $100*B$ ) =  $495/D - 450$ .#

Volume, and hence body density, can be accurately measured a variety of ways. The technique of underwater weighing "computes body volume as the difference between body weight measured in air and weight measured during water submersion. In other words, body volume is equal to the loss of weight in water with the appropriate temperature correction for the water's density" (Katch and McArdle (1977), p. 113). Using this technique,

Body Density =  $WA/[(WA-WW)/c.f. - LV]$

where  $WA$  = Weight in air (kg)  $WW$  = Weight in water (kg)  $c.f.$  = Water correction factor (=1 at 39.2 deg F as one-gram of water occupies exactly one  $\text{cm}^3$  at this temperature, =.997 at 76-78 deg F)  $LV$  = Residual Lung Volume (liters)

(Katch and McArdle (1977), p. 115). Other methods of determining body volume are given in Behnke and Wilmore (1974), p. 22 ff.

Measurement standards are apparently those listed in Behnke and Wilmore (1974), pp. 45-48 where, for instance, the abdomen circumference is measured "laterally, at the level of the iliac crests, and anteriorly, at the umbilicus".)

## Source

These data are used to produce the predictive equations for lean body weight given in the abstract "Generalized body composition prediction equation for men using simple measurement techniques", K.W. Penrose, A.G. Nelson, A.G. Fisher, FACSM, Human Performance Research Center, Brigham Young University, Provo, Utah 84602 as listed in *Medicine and Science in Sports and Exercise*, vol. 17, no. 2, April 1985, p. 189. (The predictive equations were obtained from the first 143 of the 252 cases that are listed below). The data were generously supplied by Dr. A. Garth Fisher who gave permission to freely distribute the data and use for non-commercial purposes.

## References

Bailey, Covert (1994). *Smart Exercise: Burning Fat, Getting Fit*, Houghton-Mifflin Co., Boston, pp. 179-186.

Behnke, A.R. and Wilmore, J.H. (1974). Evaluation and Regulation of Body Build and Composition, Prentice-Hall, Englewood Cliffs, N.J.

Siri, W.E. (1956), "Gross composition of the body", in Advances in Biological and Medical Physics, vol. IV, edited by J.H. Lawrence and C.A. Tobias, Academic Press, Inc., New York.

Katch, Frank and McArdle, William (1977). Nutrition, Weight Control, and Exercise, Houghton Mifflin Co., Boston.

Wilmore, Jack (1976). Athletic Training and Physical Fitness: Physiological Principles of the Conditioning Process, Allyn and Bacon, Inc., Boston.

## Examples

```
data(bodyfat)
bodyfat.bas = bas.lm(Bodyfat ~ Abdomen, data=bodyfat, prior="ZS-null")
summary(bodyfat.bas)
plot(Bodyfat ~ Abdomen, data=bodyfat, xlab="abdomen circumference (cm)")
betas = coef(bodyfat.bas)$postmean # current version has that intercept is ybar
betas[1] = betas[1] - betas[2]*bodyfat.bas$mean.x
abline(betas)
abline(coef(lm(Bodyfat ~ Abdomen, data=bodyfat)), col=2, lty=2)
```

---

CCH

*Generalized g-Prior Distribution for Coefficients in BMA Models*

---

## Description

Creates an object representing the CCH mixture of g-priors on coefficients for BAS .

## Usage

```
CCH(alpha, beta, s = 0)
```

## Arguments

alpha	a scalar $> 0$ , recommended $\alpha = .5$ (betaprime) or 1 for CCH. The hyper.g(alpha) is equivalent to CCH(alpha -2, 2, 0). Liang et al recommended values in the range $2 < \alpha_h \leq 4$
beta	a scalar $> 0$ . The value is not updated by the data; beta should be a function of n for consistency under the null model. The hyper-g corresponds to $b = 2$
s	a scalar, recommended $s=0$

## Details

Creates a structure used for [bas.glm](#).

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise A Clyde

**See Also**

[IC.prior](#), [bic.prior](#), [bas.glm](#)

Other beta priors: [EB.local\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [beta.prime\(\)](#), [g.prior\(\)](#), [hyper.g.n\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [tCCH\(\)](#), [testBF.prior\(\)](#)

**Examples**

```
CCH(alpha = .5, beta = 100, s = 0)
```

---

climate

*Climate Data*

---

**Description**

Climate Data

**Format**

Scientists are interested in the Earth's temperature change since the last glacial maximum, about 20,000 years ago. The first study to estimate the temperature change was published in 1980, and estimated a change of -1.5 degrees C, +/- 1.2 degrees C in tropical sea surface temperatures. The negative value means that the Earth was colder then than now. Since 1980 there have been many other studies. `climate` is a dataset with 63 measurements on 5 variables:

***deltaT*** the response variables, which is the change in temperature in degrees Celsius;

***sdev*** a standard deviation for the calculated *deltaT*;

***proxy*** a number 1-8 reflecting which type of measurement system was used to derive *deltaT*. Some proxies can be used over land, others over water. The proxies are coded as

- 1 "Mg/Ca"
- 2 "alkenone"
- 3 "Faunal"
- 4 "Sr/Ca"
- 5 "del 180"
- 6 "Ice Core"
- 7 "Pollen"
- 8 "Noble Gas"

***T/M*** , an indicator of whether it was a terrestrial or marine study (T/M), which is coded as 0 for Terrestrial, 1 for Marine;

***latitude*** the latitude where the data were collected.

**Source**

Data provided originally by Michael Lavine

---

 coef.bas

---

*Coefficients of a Bayesian Model Average object*


---

**Description**

Extract conditional posterior means and standard deviations, marginal posterior means and standard deviations, posterior probabilities, and marginal inclusions probabilities under Bayesian Model Averaging from an object of class 'bas'

**Usage**

```
## S3 method for class 'bas'
coef(object, n.models, estimator = "BMA", ...)

## S3 method for class 'coef.bas'
print(x, digits = max(3, getOption("digits") - 3), ...)
```

**Arguments**

object	object of class 'bas' created by BAS
n.models	Number of top models to report in the printed summary, for coef the default is to use all models. To extract summaries for the Highest Probability Model, use n.models=1 or estimator="HPM".
estimator	return summaries for a selected model, rather than using BMA. Options are 'HPM' (highest posterior probability model), 'MPM' (median probability model), and 'BMA'
...	other optional arguments
x	object of class 'coef.bas' to print
digits	number of significant digits to print

**Details**

Calculates posterior means and (approximate) standard deviations of the regression coefficients under Bayesian Model averaging using g-priors and mixtures of g-priors. Print returns overall summaries. For fully Bayesian methods that place a prior on g, the posterior standard deviations do not take into account full uncertainty regarding g. Will be updated in future releases.

**Value**

coefficients returns an object of class coef.bas with the following:

conditionalmeans	a matrix with conditional posterior means for each model
conditionalsd	standard deviations for each model
postmean	marginal posterior means of each regression coefficient using BMA
postsd	marginal posterior standard deviations using BMA
postne0	vector of posterior inclusion probabilities, marginal probability that a coefficient is non-zero

**Note**

With highly correlated variables, marginal summaries may not be representative of the joint distribution. Use `plot.coef.bas` to view distributions. The value reported for the intercept is under the centered parameterization. Under the Gaussian error model it will be centered at the sample mean of Y.

**Author(s)**

Merlise Clyde <clyde@duke.edu>

**References**

Liang, F., Paulo, R., Molina, G., Clyde, M. and Berger, J.O. (2005) Mixtures of g-priors for Bayesian Variable Selection. Journal of the American Statistical Association. 103:410-423.  
doi: [10.1198/016214507000001337](https://doi.org/10.1198/016214507000001337)

**See Also**

`bas`, `confint.coef.bas`

Other bas methods: `BAS`, `bas.lm()`, `confint.coef.bas()`, `confint.pred.bas()`, `diagnostics()`, `fitted.bas()`, `force.heredity.bas()`, `image.bas()`, `plot.confint.bas()`, `predict.basglm()`, `predict.bas()`, `summary.bas()`, `update.bas()`, `variable.names.pred.bas()`

**Examples**

```
data("Hald")
hald.gprior = bas.lm(Y~ ., data=Hald, n.models=2^4, alpha=13,
                    prior="ZS-null", initprobs="Uniform", update=10)
coef.hald.gprior = coefficients(hald.gprior)
coef.hald.gprior
plot(coef.hald.gprior)
confint(coef.hald.gprior)

#Estimation under Median Probability Model
coef.hald.gprior = coefficients(hald.gprior, estimator="MPM")
coef.hald.gprior
```

```

plot(coef.hald.gprior)
plot(confint(coef.hald.gprior))

coef.hald.gprior = coefficients(hald.gprior, estimator="HPM")
coef.hald.gprior
plot(coef.hald.gprior)
confint(coef.hald.gprior)

# To add estimation under Best Predictive Model

```

---

```

confint.coef.bas      Compute Credible Intervals for BAS regression coefficients from BAS
                      objects

```

---

### Description

Uses Monte Carlo simulations using posterior means and standard deviations of coefficients to generate draws from the posterior distributions and returns highest posterior density (HPD) credible intervals. If the number of models equals one, then use the t distribution to find intervals. These currently condition on the estimate of  $g$ . than the description above ~~

### Usage

```

## S3 method for class 'coef.bas'
confint(object, parm, level = 0.95, nsim = 10000, ...)

```

### Arguments

object	a coef.bas object
parm	a specification of which parameters are to be given credible intervals, either a vector of numbers or a vector of names. If missing, all parameters are considered.
level	the probability coverage required
nsim	number of Monte Carlo draws from the posterior distribution. Used when number of models is greater than 1.
...	other arguments to passed; none currently

### Value

A matrix (or vector) with columns giving lower and upper HPD credible limits for each parameter. These will be labeled as  $1 - \text{level}/2$  and  $1 - (1 - \text{level})/2$  in percent (by default 2.5 and 97.5).

**Note**

For mixture of g-priors these are approximate. This uses Monte Carlo sampling so results may be subject to Monte Carlo variation and larger values of nsim may be needed to reduce variability.

**Author(s)**

Merlise A Clyde

**See Also**

Other CI methods: [confint.pred.bas\(\)](#), [plot.confint.bas\(\)](#)

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

**Examples**

```
data("Hald")
hald_gprior <- bas.lm(Y~ ., data=Hald, alpha=13,
                    prior="g-prior")
coef_hald <- coef(hald_gprior)
confint(coef_hald)
confint(coef_hald, approx=FALSE, nsim=5000)
# extract just the coefficient of X4
confint(coef_hald, parm="X4")
```

---

confint.pred.bas	<i>Compute Credible (Bayesian Confidence) Intervals for a BAS predict object</i>
------------------	--

---

**Description**

Compute credible intervals for in-sample or out of sample prediction or for the regression function

**Usage**

```
## S3 method for class 'pred.bas'
confint(object, parm, level = 0.95, nsim = 10000, ...)
```

**Arguments**

object	an object created by <a href="#">predict.bas</a>
parm	character variable, "mean" or "pred". If missing parm='pred'.
level	the nominal level of the (point-wise) credible interval
nsim	number of Monte Carlo simulations for sampling methods with BMA
...	optional arguments to pass on to next function call; none at this time.

**Details**

This constructs approximate 95 percent Highest Posterior Density intervals for 'pred.bas' objects. If the estimator is based on model selection, the intervals use a Student t distribution using the estimate of  $g$ . If the estimator is based on BMA, then nsim draws from the mixture of Student t distributions are obtained with the HPD interval obtained from the Monte Carlo draws.

**Value**

a matrix with lower and upper level \* 100 percent credible intervals for either the mean of the regression function or predicted values.

**Author(s)**

Merlise A Clyde

**See Also**

[predict.bas](#)

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

Other CI methods: [confint.coef.bas\(\)](#), [plot.confint.bas\(\)](#)

**Examples**

```
data("Hald")
hald.gprior = bas.lm(Y~ ., data=Hald, alpha=13, prior="g-prior")
hald.pred = predict(hald.gprior, estimator="BPM", predict=FALSE, se.fit=TRUE)
confint(hald.pred, parm="mean")
confint(hald.pred) #default
hald.pred = predict(hald.gprior, estimator="BMA", predict=FALSE, se.fit=TRUE)
confint(hald.pred)
```



**Description**

Compute average prediction error from out of sample predictions

**Usage**

```
cv.summary.bas(pred, ytrue, score = "squared-error")
```

**Arguments**

pred	fitted or predicted value from the output from <a href="#">predict.bas</a>
ytrue	vector of left out response values
score	function used to summarize error rate. Either "squared-error", or "miss-class"

**Value**

For squared error, the average prediction error for the Bayesian estimator error =  $\sqrt{\text{sum}(y_{\text{true}} - \hat{y})^2 / n_{\text{pred}}}$  while for binary data the misclassification rate is more appropriate.

**Author(s)**

Merlise Clyde <clyde@duke.edu>

**See Also**

[predict.bas](#)

**Examples**

```
## Not run:
library(foreign)
cognitive <- read.dta("https://www.stat.columbia.edu/~gelman/arm/examples/child.iq/kidiq.dta")
cognitive$mom_work <- as.numeric(cognitive$mom_work > 1)
cognitive$mom_hs <- as.numeric(cognitive$mom_hs > 0)
colnames(cognitive) <- c("kid_score", "hs", "iq", "work", "age")

set.seed(42)
n <- nrow(cognitive)
test <- sample(1:n, size = round(.20 * n), replace = FALSE)
testdata <- cognitive[test, ]
traindata <- cognitive[-test, ]
cog_train <- bas.lm(kid_score ~ ., prior = "BIC", modelprior = uniform(), data = traindata)
yhat <- predict(cog_train, newdata = testdata, estimator = "BMA", se = F)
cv.summary.bas(yhat$fit, testdata$kid_score)

## End(Not run)
```

---

diagnostics

*BAS MCMC diagnostic plot*

---

### Description

Function to help assess convergence of MCMC sampling for bas objects.

### Usage

```
diagnostics(obj, type = c("pip", "model"), ...)
```

### Arguments

obj	an object created by <code>bas.lm</code> or <code>bas.glm</code>
type	type of diagnostic plot. If "pip" the marginal inclusion probabilities are used, while if "model", plot posterior model probabilities
...	additional graphics parameters to be passed to plot

### Details

BAS calculates posterior model probabilities in two ways when `method="MCMC"`. The first is using the relative Monte Carlo frequencies of sampled models. The second is to renormalize the marginal likelihood times prior probabilities over the sampled models. If the Markov chain has converged, these two quantities should be the same and fall on a 1-1 line. If not, running longer may be required. If the chain has not converged, the Monte Carlo frequencies may have less bias, although may exhibit more variability on repeated runs.

### Value

a plot with of the marginal inclusion probabilities (pip) estimated by MCMC and renormalized marginal likelihoods times prior probabilities or model probabilities.

### Author(s)

Merlise Clyde (<clyde@duke.edu>)

### See Also

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

## Examples

```
library(MASS)
data(UScrime)
UScrime[, -2] <- log(UScrime[, -2])
crime.ZS <- bas.lm(y ~ .,
  data = UScrime,
  prior = "ZS-null",
  modelprior = uniform(),
  method = "MCMC",
  MCMC.iter = 1000
) # short run for the example
diagnostics(crime.ZS)
```

---

EB.global

*Find the global Empirical Bayes estimates for BMA*

---

## Description

Finds the global Empirical Bayes estimates of  $g$  in Zellner's  $g$ -prior and model probabilities

## Usage

```
EB.global(object, tol = 0.1, g.0 = NULL, max.iterations = 100)
```

## Arguments

<code>object</code>	A 'bas' object created by <code>bas</code>
<code>tol</code>	tolerance for estimating $g$
<code>g.0</code>	initial value for $g$
<code>max.iterations</code>	Maximum number of iterations for the EM algorithm

## Details

Uses the EM algorithm in Liang et al to estimate the type II MLE of  $g$  in Zellner's  $g$  prior

## Value

An object of class 'bas' using Zellner's  $g$  prior with an estimate of  $g$  based on all models

## Author(s)

Merlise Clyde <clyde@stat.duke.edu>

## References

Liang, F., Paulo, R., Molina, G., Clyde, M. and Berger, J.O. (2008) Mixtures of  $g$ -priors for Bayesian Variable Selection. *Journal of the American Statistical Association*. 103:410-423.  
doi: [10.1198/016214507000001337](https://doi.org/10.1198/016214507000001337)

**See Also**

[bas](#), [update](#)

**Examples**

```
library(MASS)
data(UScrime)
UScrime[,-2] = log(UScrime[,-2])
# EB local uses a different g within each model
crime.EBL = bas.lm(y ~ ., data=UScrime, n.models=2^15,
                  prior="EB-local", initprobs= "eplogp")
# use a common (global) estimate of g
crime.EBG = EB.global(crime.EBL)
```

---

EB.local

*Empirical Bayes Prior Distribution for Coefficients in BMA Model*

---

**Description**

Creates an object representing the EB prior for BAS GLM.

**Usage**

```
EB.local()
```

**Details**

Creates a structure used for [bas.glm](#).

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[CCH](#) and [bas.glm](#)

Other beta priors: [CCH\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [beta.prior\(\)](#), [g.prior\(\)](#), [hyper.g.n\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [tCCH\(\)](#), [testBF.prior\(\)](#)

**Examples**

```
EB.local()
```

---

eplogprob	<i>eplogprob - Compute approximate marginal inclusion probabilities from pvalues</i>
-----------	--

---

### Description

eplogprob calculates approximate marginal posterior inclusion probabilities from p-values computed from a linear model using a lower bound approximation to Bayes factors. Used to obtain initial inclusion probabilities for sampling using Bayesian Adaptive Sampling `bas.lm`

### Usage

```
eplogprob(lm.obj, thresh = 0.5, max = 0.99, int = TRUE)
```

### Arguments

lm.obj	a linear model object
thresh	the value of the inclusion probability when if the p-value > 1/exp(1), where the lower bound approximation is not valid.
max	maximum value of the inclusion probability; used for the <code>bas.lm</code> function to keep initial inclusion probabilities away from 1.
int	If the Intercept is included in the linear model, set the marginal inclusion probability corresponding to the intercept to 1

### Details

Sellke, Bayarri and Berger (2001) provide a simple calibration of p-values

$$BF(p) = -e p \log(p)$$

which provide a lower bound to a Bayes factor for comparing  $H_0: \beta = 0$  versus  $H_1: \beta \neq 0$ , when the p-value  $p$  is less than  $1/e$ . Using equal prior odds on the hypotheses  $H_0$  and  $H_1$ , the approximate marginal posterior inclusion probability

$$p(\beta \neq 0 | \text{data}) = 1/(1 + BF(p))$$

When  $p > 1/e$ , we set the marginal inclusion probability to 0.5 or the value given by `thresh`.

### Value

eplogprob returns a vector of marginal posterior inclusion probabilities for each of the variables in the linear model. If `int = TRUE`, then the inclusion probability for the intercept is set to 1. If the model is not full rank, variables that are linearly dependent base on the QR factorization will have NA for their p-values. In `bas.lm`, where the probabilities are used for sampling, the inclusion probability is set to 0.

### Author(s)

Merlise Clyde <clyde@stat.duke.edu>

**References**

Sellke, Thomas, Bayarri, M. J., and Berger, James O. (2001), “Calibration of p-values for testing precise null hypotheses”, *The American Statistician*, 55, 62-71.

**See Also**

[bas](#)

**Examples**

```
library(MASS)
data(UScrime)
UScrime[, -2] = log(UScrime[, -2])
eplogprob(lm(y ~ ., data=UScrime))
```

---

eplogprob.marg	<i>eplogprob.marg - Compute approximate marginal inclusion probabilities from pvalues</i>
----------------	---

---

**Description**

eplogprob.marg calculates approximate marginal posterior inclusion probabilities from p-values computed from a series of simple linear regression models using a lower bound approximation to Bayes factors. Used to order variables and if appropriate obtain initial inclusion probabilities for sampling using Bayesian Adaptive Sampling `bas.lm`

**Usage**

```
eplogprob.marg(Y, X, thresh = 0.5, max = 0.99, int = TRUE)
```

**Arguments**

Y	response variable
X	design matrix with a column of ones for the intercept
thresh	the value of the inclusion probability when if the p-value > 1/exp(1), where the lower bound approximation is not valid.
max	maximum value of the inclusion probability; used for the <code>bas.lm</code> function to keep initial inclusion probabilities away from 1.
int	If the Intercept is included in the linear model, set the marginal inclusion probability corresponding to the intercept to 1

**Details**

Sellke, Bayarri and Berger (2001) provide a simple calibration of p-values

$$BF(p) = -e p \log(p)$$

which provide a lower bound to a Bayes factor for comparing  $H_0: \beta = 0$  versus  $H_1: \beta \neq 0$ , when the p-value  $p$  is less than  $1/e$ . Using equal prior odds on the hypotheses  $H_0$  and  $H_1$ , the approximate marginal posterior inclusion probability

$$p(\beta \neq 0 \mid \text{data}) = 1/(1 + BF(p))$$

When  $p > 1/e$ , we set the marginal inclusion probability to 0.5 or the value given by `thresh`. For the `eplogprob.marg` the marginal p-values are obtained using statistics from the  $p$  simple linear regressions

$P(F > (n-2) R^2/(1 - R^2))$  where  $F \sim F(1, n-2)$  where  $R^2$  is the square of the correlation coefficient between  $y$  and  $X_j$ .

**Value**

`eplogprob.prob` returns a vector of marginal posterior inclusion probabilities for each of the variables in the linear model. If `int = TRUE`, then the inclusion probability for the intercept is set to 1.

**Author(s)**

Merlise Clyde <clyde@stat.duke.edu>

**References**

Sellke, Thomas, Bayarri, M. J., and Berger, James O. (2001), "Calibration of p-values for testing precise null hypotheses", *The American Statistician*, 55, 62-71.

**See Also**

[bas](#)

**Examples**

```
library(MASS)
data(UScrime)
UScrime[,-2] = log(UScrime[,-2])
eplogprob(lm(y ~ ., data=UScrime))
```

fitted.bas

*Fitted values for a BAS BMA objects***Description**

Calculate fitted values for a BAS BMA object

**Usage**

```
## S3 method for class 'bas'
fitted(
  object,
  type = "link",
  estimator = "BMA",
  top = NULL,
  na.action = na.pass,
  ...
)
```

**Arguments**

object	An object of class 'bas' as created by <a href="#">bas</a>
type	type equals "response" or "link" in the case of GLMs (default is 'link')
estimator	estimator type of fitted value to return. Default is to use BMA with all models. Options include 'HPM' the highest probability model 'BMA' Bayesian model averaging, using optionally only the 'top' models 'MPM' the median probability model of Barbieri and Berger. 'BPM' the model that is closest to BMA predictions under squared error loss
top	optional argument specifying that the 'top' models will be used in constructing the BMA prediction, if NULL all models will be used. If top=1, then this is equivalent to 'HPM'
na.action	function determining what should be done with missing values in newdata. The default is to predict NA.
...	optional arguments, not used currently

**Details**

Calculates fitted values at observed design matrix using either the highest probability model, 'HPM', the posterior mean (under BMA) 'BMA', the median probability model 'MPM' or the best predictive model 'BPM'. The median probability model is defined by including variable where the marginal inclusion probability is greater than or equal to 1/2. For type="BMA", the weighted average may be based on using a subset of the highest probability models if an optional argument is given for top. By default BMA uses all sampled models, which may take a while to compute if the number of variables or number of models is large. The "BPM" is found by computing the squared distance of the vector of fitted values for a model and the fitted values under BMA and returns the



model with the smallest distance. In the presence of multicollinearity this may be quite different from the MPM, with extreme collinearity may drop relevant predictors.

### Value

A vector of length n of fitted values.

### Author(s)

Merlise Clyde <clyde@duke.edu>

### References

- Barbieri, M. and Berger, J.O. (2004) Optimal predictive model selection. *Annals of Statistics*. 32, 870-897.  
<https://projecteuclid.org/euclid.aos/1085408489?url=/UI/1.0/Summarize/euclid.aos/1085408489>
- Clyde, M. Ghosh, J. and Littman, M. (2010) Bayesian Adaptive Sampling for Variable Selection and Model Averaging. *Journal of Computational Graphics and Statistics*. 20:80-101  
 doi: [10.1198/jcgs.2010.09049](https://doi.org/10.1198/jcgs.2010.09049)

### See Also

[predict.bas](#) [predict.basglm](#)

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

Other predict methods: [predict.basglm\(\)](#), [predict.bas\(\)](#), [variable.names.pred.bas\(\)](#)

### Examples

```
data(Hald)
hald.gprior = bas.lm(Y~ ., data=Hald, prior="ZS-null", initprobs="Uniform")
plot(Hald$Y, fitted(hald.gprior, estimator="HPM"))
plot(Hald$Y, fitted(hald.gprior, estimator="BMA", top=3))
plot(Hald$Y, fitted(hald.gprior, estimator="MPM"))
plot(Hald$Y, fitted(hald.gprior, estimator="BPM"))
```

---

force.heredity.bas	<i>Post processing function to force constraints on interaction inclusion bas BMA objects</i>
--------------------	---

---

### Description

This function takes the output of a bas object and allows higher order interactions to be included only if their parent lower order interactions terms are in the model, by assigning zero prior probability, and hence posterior probability, to models that do include their respective parents.

**Usage**

```
force.heredity.bas(object, prior.prob = 0.5)
```

**Arguments**

object	a bas linear model or generalized linear model object
prior.prob	prior probability that a term is included conditional on parents being included

**Value**

a bas object with updated models, coefficients and summaries obtained removing all models with zero prior and posterior probabilities.

**Note**

Currently prior probabilities are computed using conditional Bernoulli distributions, i.e.  $P(\text{gamma}_j = 1 \mid \text{Parents}(\text{gamma}_j) = 1) = \text{prior.prob}$ . This is not very efficient for models with a large number of levels. Future updates will force this at the time of sampling.

**Author(s)**

Merlise A Clyde

**See Also**

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

**Examples**

```
data("chickwts")
bas.chk <- bas.lm(weight ~ feed, data = chickwts)
# summary(bas.chk) # 2^5 = 32 models
bas.chk.int <- force.heredity.bas(bas.chk)
# summary(bas.chk.int) # two models now

data(Hald)
bas.hald <- bas.lm(Y ~ .^2, data = Hald)
bas.hald.int <- force.heredity.bas(bas.hald)
image(bas.hald.int)

image(bas.hald.int)

# two-way interactions
data(ToothGrowth)
ToothGrowth$dose <- factor(ToothGrowth$dose)
levels(ToothGrowth$dose) <- c("Low", "Medium", "High")
TG.bas <- bas.lm(len ~ supp * dose, data = ToothGrowth, modelprior = uniform())
```

```
TG.bas.int <- force.heredity.bas(TG.bas)
image(TG.bas.int)
```

---

g.prior

*Families of G-Prior Distribution for Coefficients in BMA Models*

---

## Description

Creates an object representing the g-prior distribution on coefficients for BAS.

## Usage

```
g.prior(g)
```

## Arguments

**g** a scalar used in the covariance of Zellner's g-prior,  $\text{Cov}(\beta) = \sigma^2 g (X'X)^{-1}$

## Details

Creates a structure used for BAS.

## Value

returns an object of class "prior", with the family and hyperparameters.

## Author(s)

Merlise Clyde

## See Also

[IC.prior](#)

Other beta priors: [CCH\(\)](#), [EB.local\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [beta.prime\(\)](#), [hyper.g.n\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [tCCH\(\)](#), [testBF.prior\(\)](#)

## Examples

```
g.prior(100)
```

---

Hald

*Hald Data*


---

### Description

The Hald data have been used in many books and papers to illustrate variable selection. The data relate to an engineering application that was concerned with the effect of the composition of cement on heat evolved during hardening. The response variable  $Y$  is the *heat evolved* in a cement mix. The four explanatory variables are ingredients of the mix, X1: *tricalcium aluminate*, X2: *tricalcium silicate*, X3: *tetracalcium alumino ferrite*, X4: *dicalcium silicate*. An important feature of these data is that the variables X1 and X3 are highly correlated, as well as the variables X2 and X4. Thus we should expect any subset of (X1,X2,X3,X4) that includes one variable from highly correlated pair to do as any subset that also includes the other member.

### Format

hald is a dataframe with 13 observations and 5 variables (columns),

Y: Heat evolved per gram of cement (in calories) X1: Amount of tricalcium aluminate X2: Amount of tricalcium silicate X3: Amount of tetracalcium alumino ferrite X4: Amount of dicalcium silicate

### Source

Wood, H., Steinour, H.H., and Starke, H.R. (1932). "Effect of Composition of Portland cement on Heat Evolved During Hardening", *Industrial and Engineering Chemistry*, 24, 1207-1214.

---

hyper.g

*Hyper-g-Prior Distribution for Coefficients in BMA Models*


---

### Description

Creates an object representing the hyper-g mixture of g-priors on coefficients for BAS.

### Usage

```
hyper.g(alpha = 3)
```

### Arguments

alpha a scalar > 0. The hyper.g(alpha) is equivalent to CCH(alpha -2, 2, 0). Liang et al recommended values in the range  $2 < \alpha_h \leq 3$

### Details

Creates a structure used for `bas.glm`.

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[CCH](#) [bas.glm](#)

Other beta priors: [CCH\(\)](#), [EB.local\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [beta.prior\(\)](#), [g.prior\(\)](#), [hyper.g.n\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [tCCH\(\)](#), [testBF.prior\(\)](#)

**Examples**

```
hyper.g(alpha = 3)
```

---

hyper.g.n

*Generalized hyper-g/n Prior Distribution for g for mixtures of g-priors on Coefficients in BMA Models*

---

**Description**

Creates an object representing the hyper-g/n mixture of g-priors on coefficients for BAS. This is a special case of the tCCH prior

**Usage**

```
hyper.g.n(alpha = 3, n = NULL)
```

**Arguments**

alpha	a scalar > 0, recommended $2 < \alpha \leq 3$
n	The sample size; if NULL, the value derived from the data in the call to 'bas.glm' will be used.

**Details**

Creates a structure used for [bas.glm](#). This is a special case of the [tCCH](#), where `hyper.g.n(alpha=3, n)` is equivalent to `tCCH(alpha=1, beta=2, s=0, r=1.5, v = 1, theta=1/n)`

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[tCCH](#), [robust](#), [hyper.g](#), [CCHbas.glm](#)

Other beta priors: [CCH\(\)](#), [EB.local\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [beta.prime\(\)](#), [g.prior\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [tCCH\(\)](#), [testBF.prior\(\)](#)

**Examples**

```
n <- 500
hyper.g.n(alpha = 3, n = n)
```

---

hypergeometric1F1      *Confluent hypergeometric 1F1 function*

---

**Description**

Compute the Confluent Hypergeometric function:  ${}_1F_1(a,b,c,t) = \frac{\Gamma(b)}{\Gamma(b-a)\Gamma(a)} \int_0^1 t^{a-1} (1-t)^{b-a-1} \exp(ct) dt$

**Usage**

```
hypergeometric1F1(a, b, c, laplace = FALSE, log = TRUE)
```

**Arguments**

a	arbitrary
b	Must be greater 0
c	arbitrary
laplace	The default is to use the Cephes library; for large a or s this may return an NA, Inf or negative values., in which case you should use the Laplace approximation.
log	if TRUE, return log(1F1)

**Author(s)**

Merlise Clyde (<clyde@stat.duke.edu>)

**References**

Cephes library hyp1f1.c

**See Also**

Other special functions: [hypergeometric2F1\(\)](#), [phi1\(\)](#), [trCCH\(\)](#)

**Examples**

```
hypergeometric1F1(11.14756, 0.5, 0.00175097)
```

---

hypergeometric2F1      *Gaussian hypergeometric2F1 function*

---

**Description**

Compute the Gaussian Hypergeometric2F1 function:  ${}_2F_1(a,b,c,z) = \frac{\Gamma(b-c)}{\Gamma(b)\Gamma(c)} \int_0^1 t^{b-1} (1-t)^{c-b-1} (1-tz)^{-a} dt$

**Usage**

```
hypergeometric2F1(a, b, c, z, method = "Cephes", log = TRUE)
```

**Arguments**

a	arbitrary
b	Must be greater 0
c	Must be greater than b if $ z  < 1$ , and $c > b + a$ if $z = 1$
z	$ z  \leq 1$
method	The default is to use the Cephes library routine. This sometimes is unstable for large a or z near one returning Inf or negative values. In this case, try method="Laplace", which use a Laplace approximation for $\tau = \exp(t/(1-t))$ .
log	if TRUE, return $\log({}_2F_1)$

**Details**

The default is to use the routine hyp2f1.c from the Cephes library. If that return a negative value or Inf, one should try method="Laplace" which is based on the Laplace approximation as described in Liang et al JASA 2008. This is used in the hyper-g prior to calculate marginal likelihoods.

**Value**

if log=T returns the log of the 2F1 function; otherwise the 2F1 function.

**Author(s)**

Merlise Clyde (<clyde@duke.edu>)

**References**

Cephes library hyp2f1.c

Liang, F., Paulo, R., Molina, G., Clyde, M. and Berger, J.O. (2005) Mixtures of g-priors for Bayesian Variable Selection. Journal of the American Statistical Association. 103:410-423.  
doi: [10.1198/016214507000001337](https://doi.org/10.1198/016214507000001337)

**See Also**

Other special functions: [hypergeometric1F1\(\)](#), [phi1\(\)](#), [trCCH\(\)](#)

**Examples**

```
hypergeometric2F1(12, 1, 2, .65)
```

---

IC.prior	<i>Information Criterion Families of Prior Distribution for Coefficients in BMA Models</i>
----------	--

---

**Description**

Creates an object representing the prior distribution on coefficients for BAS.

**Usage**

```
IC.prior(penalty)
```

**Arguments**

penalty            a scalar used in the penalized loglikelihood of the form  $\text{penalty} \times \text{dimension}$

**Details**

The log marginal likelihood is approximated as  $-2 \times (\text{deviance} + \text{penalty} \times \text{dimension})$ . Allows alternatives to AIC (penalty = 2) and BIC (penalty =  $\log(n)$ ). For BIC, the argument may be missing, in which case the sample size is determined from the call to 'bas.glm' and used to determine the penalty.

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[g.prior](#)

Other beta priors: [CCH\(\)](#), [EB.local\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [beta.prime\(\)](#), [g.prior\(\)](#), [hyper.g.n\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [tCCH\(\)](#), [testBF.prior\(\)](#)

**Examples**

```
IC.prior(2)
aic.prior()
bic.prior(100)
```



---

 image.bas

---

*Images of models used in Bayesian model averaging*


---

## Description

Creates an image of the models selected using [bas](#).

## Usage

```
## S3 method for class 'bas'
image(
  x,
  top.models = 20,
  intensity = TRUE,
  prob = TRUE,
  log = TRUE,
  rotate = TRUE,
  color = "rainbow",
  subset = NULL,
  drop.always.included = FALSE,
  offset = 0.75,
  digits = 3,
  vlas = 2,
  plas = 0,
  rlas = 0,
  ...
)
```

## Arguments

x	A BMA object of type 'bas' created by BAS
top.models	Number of the top ranked models to plot
intensity	Logical variable, when TRUE image intensity is proportional to the probability or log(probability) of the model, when FALSE, intensity is binary indicating just presence (light) or absence (dark) of a variable.
prob	Logical variable for whether the area in the image for each model should be proportional to the posterior probability (or log probability) of the model (TRUE) or with equal area (FALSE).
log	Logical variable indicating whether the intensities should be based on log posterior odds (TRUE) or posterior probabilities (FALSE). The log of the posterior odds is for comparing the each model to the worst model in the top.models.
rotate	Should the image of models be rotated so that models are on the y-axis and variables are on the x-axis (TRUE)
color	The color scheme for image intensities. The value "rainbow" uses the rainbow palette. The value "blackandwhite" produces a black and white image (greyscale image)

subset	indices of variables to include/exclude in plot
drop.always.included	logical variable to drop variables that are always forced into the model. FALSE by default.
offset	numeric value to add to intensity
digits	number of digits in posterior probabilities to keep
vlas	las parameter for placing variable names; see par
plas	las parameter for posterior probability axis
rlas	las parameter for model ranks
...	Other parameters to be passed to the image and axis functions.

### Details

Creates an image of the model space sampled using [bas](#). If a subset of the top models are plotted, then probabilities are renormalized over the subset.

### Note

Suggestion to allow area of models be proportional to posterior probability due to Thomas Lumley

### Author(s)

Merlise Clyde <clyde@stat.duke.edu>

### References

Clyde, M. (1999) Bayesian Model Averaging and Model Search Strategies (with discussion). In Bayesian Statistics 6. J.M. Bernardo, A.P. Dawid, J.O. Berger, and A.F.M. Smith eds. Oxford University Press, pages 157-185.

### See Also

[bas](#)

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

Other bas plots: [plot.bas\(\)](#), [plot.coef.bas\(\)](#)

### Examples

```
require(graphics)
data("Hald")
hald.ZSprior <- bas.lm(Y ~ ., data = Hald, prior = "ZS-null")
image(hald.ZSprior, drop.always.included = TRUE) # drop the intercept
```

---

`intrinsic`*Intrinsic Prior Distribution for Coefficients in BMA Models*

---

**Description**

Creates an object representing the intrinsic prior on  $g$ , a special case of the tCCH mixture of  $g$ -priors on coefficients for BAS.

**Usage**

```
intrinsic(n = NULL)
```

**Arguments**

`n` the sample size; if `NULL`, the value derived from the data in the call to `'bas.glm'` will be used.

**Details**

Creates a structure used for `bas.glm`.

**Value**

returns an object of class "prior", with the family "intrinsic" of class "TCCH" and hyperparameters  $\alpha = 1$ ,  $\beta = 1$ ,  $s = 0$ ,  $r = 1$ ,  $n = n$  for the tCCH prior where  $\theta$  in the tCCH prior is determined by the model size and sample size.

**Author(s)**

Merlise A Clyde

**References**

Womack, A., Novelo, L.L., Casella, G. (2014). "Inference From Intrinsic Bayes' Procedures Under Model Selection and Uncertainty". *Journal of the American Statistical Association*. 109:1040-1053. <https://amstat.tandfonline.com/doi/abs/10.1080/01621459.2014.880348>

**See Also**

`tCCH`, `robust`, `hyper.g`, `hyper.g.nbas.glm`

Other beta priors: `CCH()`, `EB.local()`, `IC.prior()`, `Jeffreys()`, `TG()`, `beta.prime()`, `g.prior()`, `hyper.g.n()`, `hyper.g()`, `robust()`, `tCCH()`, `testBF.prior()`

**Examples**

```
n <- 500
tCCH(alpha = 1, beta = 2, s = 0, r = 1.5, v = 1, theta = 1 / n)
```

---

Jeffreys

*Jeffreys Prior Distribution for  $g$  for Mixtures of  $g$ -Priors for Coefficients in BMA Models*

---

### Description

Creates an object representing the Jeffrey's Prior on  $g$  mixture of  $g$ -priors on coefficients for BAS. This is equivalent to a limiting version of the  $CCH(a, 2, 0)$  with  $a = 0$  or they hyper- $g(a = 2)$  and is an improper prior. As  $g$  does not appear in the Null Model, Bayes Factors and model probabilities are not well-defined because of arbitrary normalizing constants, and for this reason the null model is excluded and the same constants are used across other models.

### Usage

`Jeffreys()`

### Details

Creates a structure used for `bas.glm`.

### Value

returns an object of class "prior", with the family and hyperparameters.

### Author(s)

Merlise Clyde

### See Also

`CCH` `bas.glm`

Other beta priors: `CCH()`, `EB.local()`, `IC.prior()`, `TG()`, `beta.prime()`, `g.prior()`, `hyper.g.n()`, `hyper.g()`, `intrinsic()`, `robust()`, `tCCH()`, `testBF.prior()`

### Examples

`Jeffreys()`

---

list2matrix.bas	<i>Coerce a BAS list object into a matrix.</i>
-----------------	--

---

### Description

Models, coefficients, and standard errors in objects of class 'bas' are represented as a list of lists to reduce storage by omitting the zero entries. These functions coerce the list object to a matrix and fill in the zeros to facilitate other computations.

### Usage

```
list2matrix.bas(x, what, which.models = NULL)
```

### Arguments

x	a 'bas' object
what	name of bas list to coerce
which.models	a vector of indices use to extract a subset

### Details

list2matrix.bas(x, which) is equivalent to list2matrix.which(x), however, the latter uses sapply rather than a loop. list2matrix.which and which.matrix both coerce x\$which into a matrix.

### Value

a matrix representation of x\$what, with number of rows equal to the length of which.models or total number of models and number of columns x\$n.vars

### Author(s)

Merlise Clyde <clyde@duke.edu>

### See Also

[bas](#)

Other as.matrix methods: [list2matrix.which\(\)](#), [which.matrix\(\)](#)

### Examples

```
data(Hald)
hald.bic <- bas.lm(Y ~ ., data=Hald, prior="BIC",
                 initprobs= "eplogp")
coef <- list2matrix.bas(hald.bic, "mle") # extract all coefficients
se <- list2matrix.bas(hald.bic, "mle.se")
models <- list2matrix.which(hald.bic) #matrix of model indicators
```

```
models <- which.matrix(hald.bic$which, hald.bic$n.vars) #matrix of model indicators
```

---

`list2matrix.which`      *Coerce a BAS list object into a matrix.*

---

### Description

Models, coefficients, and standard errors in objects of class 'bas' are represented as a list of lists to reduce storage by omitting the zero entries. These functions coerce the list object to a matrix and fill in the zeros to facilitate other computations.

### Usage

```
list2matrix.which(x, which.models = NULL)
```

### Arguments

`x`                    a 'bas' object  
`which.models`      a vector of indices use to extract a subset

### Details

`list2matrix.bas(x, which)` is equivalent to `list2matrix.which(x)`, however, the latter uses `sapply` rather than a loop. `list2matrix.which` and `which.matrix` both coerce `x$which` into a matrix.

### Value

a matrix representation of `x$what`, with number of rows equal to the length of `which.models` or total number of models and number of columns `x$n.vars`

### Author(s)

Merlise Clyde <clyde@duke.edu>

### See Also

[bas](#)

Other `as.matrix` methods: [list2matrix.bas\(\)](#), [which.matrix\(\)](#)

**Examples**

```

data(Hald)
Hald.bic <- bas.lm(Y ~ ., data=Hald, prior="BIC", initprobs="eplogp")
coef <- list2matrix.bas(Hald.bic, "mle") # extract all ols coefficients
se <- list2matrix.bas(Hald.bic, "mle.se")
models <- list2matrix.which(Hald.bic) #matrix of model indicators
models <- which.matrix(Hald.bic$which, Hald.bic$n.vars) #matrix of model indicators

```

---

phi1

*Compound Confluent hypergeometric function of two variables*


---

**Description**

Compute the Confluent Hypergeometric function of two variables, also know as a Horn hypergeometric function or Humbert's hypergeometric used in Gordy (1998) with integral representation:

**Usage**

```
phi1(a, b, c, x, y, log = FALSE)
```

**Arguments**

a	a > 0
b	arbitrary
c	c > 0
x	x > 0
y	y > 0
log	logical indicating whether to return phi1 on the log scale

**Details**

$\text{phi}_1(a,b,c,x,y) = [\Gamma(c)/\Gamma(a)\Gamma(c-a)] \int_0^1 t^{a-1} (1-t)^{c-a-1} (1-yt)^{-b} \exp(xt) dt$  [https://en.wikipedia.org/wiki/Humbert\\_series](https://en.wikipedia.org/wiki/Humbert_series) Note that Gordy's arguments for x and y are reversed in the reference above.

The original 'phi1' function in 'BAS' was based on 'C' code provided by Gordy. This function returns NA's when x is greater than  $\log(\text{Machine}\$double.xmax)/2$ . A more stable method for calculating the 'phi1' function using R's 'integrate' was suggested by Daniel Heemann and is now an option whenever  $x$  is too large. For calculating Bayes factors that use the 'phi1' function we recommend using the 'log=TRUE' option to compute log Bayes factors.

**Author(s)**

Merlise Clyde (<clyde@duke.edu>)

Daniel Heemann (<df.heemann@gmail.com>)

**References**

Gordy 1998

**See Also**

Other special functions: [hypergeometric1F1\(\)](#), [hypergeometric2F1\(\)](#), [trCCH\(\)](#)

**Examples**

```
# special cases
# phi1(a, b, c, x=0, y) is the same as 2F1(b, a; c, y)
phi1(1, 2, 1.5, 0, 1 / 100, log=FALSE)
hypergeometric2F1(2, 1, 1.5, 1 / 100, log = FALSE)

# phi1(a,0,c,x,y) is the same as 1F1(a,c,x)
phi1(1, 0, 1.5, 3, 1 / 100)
hypergeometric1F1(1, 1.5, 3, log = FALSE)

# use direct integration
phi1(1, 2, 1.5, 1000, 0, log=TRUE)
```

---

plot.bas

*Plot Diagnostics for an BAS Object*

---

**Description**

Four plots (selectable by `which`) are currently available: a plot of residuals against fitted values, Cumulative Model Probabilities, log marginal likelihoods versus model dimension, and marginal inclusion probabilities.

**Usage**

```
## S3 method for class 'bas'
plot(
  x,
  which = c(1:4),
  caption = c("Residuals vs Fitted", "Model Probabilities", "Model Complexity",
    "Inclusion Probabilities"),
  panel = if (add.smooth) panel.smooth else points,
  sub.caption = NULL,
  main = "",
  ask = prod(par("mfcol")) < length(which) && dev.interactive(),
  col.in = 2,
  col.ex = 1,
  col.pch = 1,
  cex.lab = 1,
  ...,
```



```

    id.n = 3,
    labels.id = NULL,
    cex.id = 0.75,
    add.smooth = getOption("add.smooth"),
    label.pos = c(4, 2),
    subset = NULL,
    drop.always.included = FALSE
)

```

### Arguments

x	bas BMA object result of 'bas'
which	if a subset of the plots is required, specify a subset of the numbers '1:4'
caption	captions to appear above the plots
panel	panel function. The useful alternative to 'points', 'panel.smooth' can be chosen by 'add.smooth = TRUE'
sub.caption	common title-above figures if there are multiple; used as 'sub' (s.'title') otherwise. If 'NULL', as by default, a possible shortened version of <code>deparse(x\$call)</code> is used
main	title to each plot-in addition to the above 'caption'
ask	logical; if 'TRUE', the user is asked before each plot, see 'par(ask=.)'
col.in	color for the included variables
col.ex	color for the excluded variables
col.pch	color for points in panels 1-3
cex.lab	graphics parameter to control size of variable names
...	other parameters to be passed through to plotting functions
id.n	number of points to be labeled in each plot, starting with the most extreme
labels.id	vector of labels, from which the labels for extreme points will be chosen. 'NULL' uses observation numbers
cex.id	magnification of point labels.
add.smooth	logical indicating if a smoother should be added to most plots; see also 'panel' above
label.pos	positioning of labels, for the left half and right half of the graph respectively, for plots 1-4
subset	indices of variables to include/exclude in plot of marginal posterior inclusion probabilities (NULL).
drop.always.included	logical variable to drop marginal posterior inclusion probabilities for variables that are always forced into the model. FALSE by default.

**Details**

This provides a panel of 4 plots: the first is a plot of the residuals versus fitted values under BMA. The second is a plot of the cumulative marginal likelihoods of models; if the model space cannot be enumerated then this provides some indication of whether the probabilities are leveling off. The third is a plot of log marginal likelihood versus model dimension and the fourth plot show the posterior marginal inclusion probabilities.

**Author(s)**

Merlise Clyde, based on plot.lm by John Maindonald and Martin Maechler

**See Also**

[plot.coef.bas](#) and [image.bas](#).

Other bas plots: [image.bas\(\)](#), [plot.coef.bas\(\)](#)

**Examples**

```
data(Hald)
hald.gprior = bas.lm(Y~ ., data=Hald, prior="g-prior", alpha=13,
                    modelprior=beta.binomial(1,1),
                    initprobs="eplogp")

plot(hald.gprior)
```

---

plot.coef.bas	<i>Plots the posterior distributions of coefficients derived from Bayesian model averaging</i>
---------------	--

---

**Description**

Displays plots of the posterior distributions of the coefficients generated by Bayesian model averaging over linear regression.

**Usage**

```
## S3 method for class 'coef.bas'
plot(x, e = 1e-04, subset = 1:x$n.vars, ask = TRUE, ...)
```

**Arguments**

x	object of class coef.bas
e	optional numeric value specifying the range over which the distributions are to be graphed.
subset	optional numerical vector specifying which variables to graph (including the intercept)
ask	Prompt for next plot
...	other parameters to be passed to plot and lines

**Details**

Produces plots of the posterior distributions of the coefficients under model averaging. The posterior probability that the coefficient is zero is represented by a solid line at zero, with height equal to the probability. The nonzero part of the distribution is scaled so that the maximum height is equal to the probability that the coefficient is nonzero.

The parameter e specifies the range over which the distributions are to be graphed by specifying the tail probabilities that dictate the range to plot over.

**Note**

For mixtures of g-priors, uncertainty in g is not incorporated at this time, thus results are approximate

**Author(s)**

based on function plot.bic by Ian Painter in package BMA; adapted for 'bas' class by Merlise Clyde <clyde@stat.duke.edu>

**References**

Hoeting, J.A., Raftery, A.E. and Madigan, D. (1996). A method for simultaneous variable selection and outlier identification in linear regression. *Computational Statistics and Data Analysis*, 22, 251-270.

**See Also**

[coef.bas](#)

Other bas plots: [image.bas\(\)](#), [plot.bas\(\)](#)

**Examples**

```
## Not run: library(MASS)
data(UScrime)
UScrime[,-2] <- log(UScrime[,-2])
crime_bic <- bas.lm(y ~ ., data=UScrime, n.models=2^15, prior="BIC")
plot(coefficients(crime_bic), ask=TRUE)
```

```
## End(Not run)
```

---

```
plot.confint.bas      Plot Bayesian Confidence Intervals
```

---

## Description

Function takes the the output of functions that return credible intervals from BAS objects, and creates a plot of the posterior mean with segments representing the credible interval. of what the function does. ~~

## Usage

```
## S3 method for class 'confint.bas'
plot(x, horizontal = FALSE, ...)
```

## Arguments

x	the output from <a href="#">confint.coef.bas</a> or <a href="#">confint.pred.bas</a> containing credible intervals and estimates.
horizontal	orientation of the plot
...	optional graphical arguments to pass on to plot

## Details

This function takes the HPD intervals or credible intervals created by [confint.coef.bas](#) or [confint.pred.bas](#) from BAS objects, and creates a plot of the posterior mean with segments representing the credible interval. BAS tries to return HPD intervals, and under model averaging these may not be symmetric. the description above ~~

## Value

A plot of the credible intervals.

## Author(s)

Merlise A Clyde

## See Also

[confint.coef.bas](#), [confint.pred.bas](#), [coef.bas](#), [predict.bas](#), [link{bas.lm}](#)

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

Other CI methods: [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#)

**Examples**

```

data(Hald)
hald.ZS = bas.lm(Y ~ ., data=Hald, prior="ZS-null", modelprior=uniform())
hald.coef = confint(coef(hald.ZS), parm=2:5)
plot(hald.coef)
plot(hald.coef, horizontal=TRUE)
plot(confint(predict(hald.ZS, se.fit=TRUE), parm="mean"))

```

---

predict.bas

*Prediction Method for an object of class BAS*


---

**Description**

Predictions under model averaging or other estimators from a BMA object of class inheriting from 'bas'.

**Usage**

```

## S3 method for class 'bas'
predict(
  object,
  newdata,
  se.fit = FALSE,
  type = "link",
  top = NULL,
  estimator = "BMA",
  na.action = na.pass,
  ...
)

```

**Arguments**

object	An object of class BAS, created by bas
newdata	dataframe for predictions. If missing, then use the dataframe used for fitting for obtaining fitted and predicted values.
se.fit	indicator for whether to compute se of fitted and predicted values
type	Type of predictions required. "link" which is on the scale of the linear predictor is the only option currently for linear models, which for the normal model is equivalent to type='response'.
top	a scalar integer M. If supplied, subset the top M models, based on posterior probabilities for model predictions and BMA.

estimator	estimator used for predictions. Currently supported options include: 'HPM' the highest probability model 'BMA' Bayesian model averaging, using optionally only the 'top' models 'MPM' the median probability model of Barbieri and Berger. 'BPM' the model that is closest to BMA predictions under squared error loss. BMA may be computed using only the 'top' models if supplied
na.action	function determining what should be done with missing values in newdata. The default is to predict NA.
...	optional extra arguments

### Details

Use BMA and/or model selection to form predictions using the top highest probability models.

### Value

a list of	
fit	fitted values based on the selected estimator
Ybma	predictions using BMA, the same as fit for non-BMA methods for compatibility; will be deprecated
Ypred	matrix of predictions under each model for BMA
se.fit	se of fitted values; in the case of BMA this will be a matrix
se.pred	se for predicted values; in the case of BMA this will be a matrix
se.bma.fit	vector of posterior sd under BMA for posterior mean of the regression function. This will be NULL if estimator is not 'BMA'
se.bma.pred	vector of posterior sd under BMA for posterior predictive values. this will be NULL if estimator is not 'BMA'
best	index of top models included
bestmodels	subset of bestmodels used for fitting or prediction
best.vars	names of variables in the top model; NULL if estimator='BMA'
df	scalar or vector of degrees of freedom for models
estimator	estimator upon which 'fit' is based.

### Author(s)

Merlise Clyde

### See Also

[bas](#), [fitted.bas](#), [confint.pred.bas](#), [variable.names.pred.bas](#)

Other predict methods: [fitted.bas\(\)](#), [predict.basglm\(\)](#), [variable.names.pred.bas\(\)](#)

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

**Examples**

```

data("Hald")
hald.gprior = bas.lm(Y ~ ., data=Hald, alpha=13, prior="g-prior")

predict(hald.gprior, newdata=Hald, estimator="BPM", se.fit=TRUE)
# same as fitted
fitted(hald.gprior, estimator="BPM")
# default is BMA and estimation of mean vector
hald.bma = predict(hald.gprior, top=5, se.fit=TRUE)
confint(hald.bma)

hald.bpm = predict(hald.gprior, newdata=Hald[1,],
                  se.fit=TRUE,
                  estimator="BPM")
confint(hald.bpm)
# extract variables
variable.names(hald.bpm)

hald.hpm = predict(hald.gprior, newdata=Hald[1,],
                  se.fit=TRUE,
                  estimator="HPM")
confint(hald.hpm)
variable.names(hald.hpm)

hald.mpm = predict(hald.gprior, newdata=Hald[1,],
                  se.fit=TRUE,
                  estimator="MPM")
confint(hald.mpm)
variable.names(hald.mpm)

```

---

predict.basglm

*Prediction Method for an Object of Class basglm*


---

**Description**

Predictions under model averaging from a BMA (BAS) object for GLMs under different loss functions.

**Usage**

```

## S3 method for class 'basglm'
predict(
  object,
  newdata,
  se.fit = FALSE,
  type = c("response", "link"),
  top = NULL,

```

```

    estimator = "BMA",
    na.action = na.pass,
    ...
)

```

### Arguments

object	An object of class "basglm", created by <code>bas.glm</code>
newdata	dataframe, new matrix or vector of data for predictions. May include a column for the intercept or just the predictor variables. If a dataframe, the variables are extracted using <code>model.matrix</code> using the call that created 'object'. May be missing in which case the data used for fitting will be used for prediction.
se.fit	indicator for whether to compute se of fitted and predicted values
type	Type of predictions required. The default is "response" is on the scale of the response variable, with the alternative being on the linear predictor scale, 'type = 'link''. Thus for a default binomial model 'type = 'response'' gives the predicted probabilities, while with 'link', the estimates are of log-odds (probabilities on logit scale).
top	A scalar integer M. If supplied, calculate results using the subset of the top M models based on posterior probabilities.
estimator	estimator used for predictions. Currently supported options include: 'HPM' the highest probability model 'BMA' Bayesian model averaging, using optionally only the 'top' models 'MPM' the median probability model of Barbieri and Berger. 'BPM' the model that is closest to BMA predictions under squared error loss. BMA may be computed using only the 'top' models if supplied
na.action	function determining what should be done with missing values in newdata. The default is to predict NA.
...	optional extra arguments

### Details

This function first calls the `predict` method for class `bas` (linear models) to form predictions on the linear predictor scale for 'BMA', 'HPM', 'MPM' etc. If the estimator is 'BMA' and 'type='response', then the inverse link is applied to fitted values for type equal 'link' and model averaging takes place in the 'response' scale. Thus applying the inverse link to BMA estimate with 'type = 'link' is not equal to the fitted values for 'type = 'response' under BMA due to the nonlinear transformation under the inverse link.

### Value

a list of	
fit	predictions using BMA or other estimators
Ypred	matrix of predictions under model(s)
postprobs	renormalized probabilities of the top models
best	index of top models included



**Author(s)**

Merlise Clyde

**See Also**[bas.glm](#), [predict.bas](#), [fitted.bas](#)Other predict methods: [fitted.bas\(\)](#), [predict.bas\(\)](#), [variable.names.pred.bas\(\)](#)Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)**Examples**

```

data(Pima.tr, package="MASS")
data(Pima.te, package="MASS")
Pima.bas = bas.glm(type ~ ., data=Pima.tr, n.models= 2^7, method="BAS",
                  betaprior=CCH(a=1, b=nrow(Pima.tr)/2, s=0), family=binomial(),
                  modelprior=uniform())
pred = predict(Pima.bas, newdata=Pima.te, top=1) # Highest Probability model
cv.summary.bas(pred$fit, Pima.te$type, score="miss-class")

```

---

print.bas

---

*Print a Summary of Bayesian Model Averaging objects from BAS*


---

**Description**

summary and print methods for Bayesian model averaging objects created by bas Bayesian Adaptive Sampling

**Usage**

```

## S3 method for class 'bas'
print(x, digits = max(3L, getOption("digits") - 3L), ...)

```

**Arguments**

x	object of class 'bas'
digits	optional number specifying the number of digits to display
...	other parameters to be passed to print.default

**Details**

The print methods display a view similar to `print.lm`. The summary methods display a view specific to Bayesian model averaging giving the top 5 highest probability models represented by their inclusion indicators. Summaries of the models include the Bayes Factor (BF) of each model to the model with the largest marginal likelihood, the posterior probability of the models, R2, dim (which includes the intercept) and the log of the marginal likelihood.

**Author(s)**

Merlise Clyde <clyde@stat.duke.edu>

**See Also**

[coef.bas](#)

**Examples**

```
library(MASS)
data(UScrime)
UScrime[, -2] <- log(UScrime[, -2])
crime.bic <- bas.lm(y ~ ., data = UScrime, n.models = 2^15, prior = "BIC", initprobs = "eplogp")
print(crime.bic)
summary(crime.bic)
```

---

protein

*Protein Activity Data*

---

**Description**

This data sets includes several predictors of protein activity from an experiment run at Glaxo.

**Format**

protein is a dataframe with 96 observations and 8 predictor variables of protein activity:

[,1]	buf	factor	Buffer
[,2]	pH	numeric	
[,3]	NaCl	numeric	
[,4]	con	numeric	protein concentration
[,5]	ra	factor	reducing agent
[,6]	det	factor	detergent
[,7]	MgCl2	numeric	
[,8]	temp	numeric	(temperature)
[,9]	prot.act1	numeric	
[,10]	prot.act2	numeric	
[,11]	prot.act3	numeric	
[,12]	prot.act4	numeric	protein activity

**Source**

Clyde, M. A. and Parmigiani, G. (1998), Protein Construct Storage: Bayesian Variable Selection and Prediction with Mixtures, *Journal of Biopharmaceutical Statistics*, 8, 431-443

---

robust

*Robust-Prior Distribution for Coefficients in BMA Model*

---

**Description**

Creates an object representing the robust prior of Bayarri et al (2012) that is mixture of g-priors on coefficients for BAS.

**Usage**

```
robust(n = NULL)
```

**Arguments**

n                    the sample size.

**Details**

Creates a prior structure used for [bas.glm](#).

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[CCH](#) and [bas.glm](#)

Other beta priors: [CCH\(\)](#), [EB.local\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [beta.prime\(\)](#), [g.prior\(\)](#), [hyper.g.n\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [tCCH\(\)](#), [testBF.prior\(\)](#)

**Examples**

```
robust(100)
```

summary.bas

*Summaries of Bayesian Model Averaging objects from BAS***Description**

summary and print methods for Bayesian model averaging objects created by bas Bayesian Adaptive Sampling

**Usage**

```
## S3 method for class 'bas'
summary(object, n.models = 5, ...)
```

**Arguments**

object	object of class 'bas'
n.models	optional number specifying the number of best models to display in summary
...	other parameters to be passed to summary.default

**Details**

The print methods display a view similar to print.lm . The summary methods display a view specific to Bayesian model averaging giving the top 5 highest probability models represented by their inclusion indicators. Summaries of the models include the Bayes Factor (BF) of each model to the model with the largest marginal likelihood, the posterior probability of the models, R2, dim (which includes the intercept) and the log of the marginal likelihood.

**Author(s)**

Merlise Clyde <clyde@duke.edu>

**See Also**

[coef.bas](#)

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [update.bas\(\)](#), [variable.names.pred.bas\(\)](#)

**Examples**

```
data(UScrime, package = "MASS")
UScrime[, -2] <- log(UScrime[, -2])
crime.bic <- bas.lm(y ~ ., data = UScrime, n.models = 2^15, prior = "BIC", initprobs = "eplogp")
print(crime.bic)
summary(crime.bic)
```

---

tCCH	<i>Generalized tCCH g-Prior Distribution for Coefficients in BMA Models</i>
------	---

---

**Description**

Creates an object representing the tCCH mixture of g-priors on coefficients for BAS.

**Usage**

```
tCCH(alpha = 1, beta = 2, s = 0, r = 3/2, v = 1, theta = 1)
```

**Arguments**

alpha	a scalar > 0, recommended alpha=.5 (betaprime) or 1.
beta	a scalar > 0. The value is not updated by the data; beta should be a function of n for consistency under the null model.
s	a scalar, recommended s=0 a priori
r	r arbitrary; in the hyper-g-n prior sets $r = (\alpha + 2)$
v	$0 < v$
theta	$\theta > 1$

**Details**

Creates a structure used for [bas.glm](#).

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[CCH](#), [robust](#), [hyper.g](#), [hyper.g.nbas.glm](#)

Other beta priors: [CCH\(\)](#), [EB.local\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [beta.primo\(\)](#), [g.prior\(\)](#), [hyper.g.n\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [testBF.prior\(\)](#)

**Examples**

```
n <- 500
tCCH(alpha = 1, beta = 2, s = 0, r = 1.5, v = 1, theta = 1 / n)
```

testBF.prior

*Test based Bayes Factors for BMA Models***Description**

Creates an object representing the prior distribution on coefficients for BAS that corresponds to the test-based Bayes Factors.

**Usage**

```
testBF.prior(g)
```

**Arguments**

**g** a scalar used in the covariance of Zellner's g-prior,  $\text{Cov}(\beta) = \sigma^2 g (X'X)^{-1}$

**Details**

Creates a prior object structure used for BAS in 'bas.glm'.

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[g.prior](#), [bas.glm](#)

Other beta priors: [CCH\(\)](#), [EB.local\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [TG\(\)](#), [beta.prime\(\)](#), [g.prior\(\)](#), [hyper.g.n\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [tCCH\(\)](#)

**Examples**

```
testBF.prior(100)
library(MASS)
data(Pima.tr)

# use g = n
bas.glm(type ~ .,
  data = Pima.tr, family = binomial(),
  betaprior = testBF.prior(nrow(Pima.tr)),
  modelprior = uniform(), method = "BAS"
)
```

**Description**

Creates an object representing the Truncated Gamma (tCCH) mixture of g-priors on coefficients for BAS, where  $u = 1/(1+g)$  has a Gamma distribution supported on  $(0, 1]$ .

**Usage**

```
TG(alpha = 2)
```

**Arguments**

alpha            a scalar  $> 0$ , recommended alpha=.5 (betaprime) or 1. alpha=2 corresponds to the uniform prior on the shrinkage factor.

**Details**

Creates a structure used for [bas.glm](#).

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[CCH](#) [bas.glm](#)

Other beta priors: [CCH\(\)](#), [EB.local\(\)](#), [IC.prior\(\)](#), [Jeffreys\(\)](#), [beta.primo\(\)](#), [g.prior\(\)](#), [hyper.g.n\(\)](#), [hyper.g\(\)](#), [intrinsic\(\)](#), [robust\(\)](#), [tCCH\(\)](#), [testBF.prior\(\)](#)

**Examples**

```
TG(alpha = 2)
CCH(alpha = 2, beta = 100, s = 0)
```

---

tr.beta.binomial	<i>Truncated Beta-Binomial Prior Distribution for Models</i>
------------------	--

---

### Description

Creates an object representing the prior distribution on models for BAS using a truncated Beta-Binomial Distribution on the Model Size

### Usage

```
tr.beta.binomial(alpha = 1, beta = 1, trunc)
```

### Arguments

alpha	parameter in the beta prior distribution
beta	parameter in the beta prior distribution
trunc	parameter that determines truncation in the distribution i.e. $P(M; \alpha, \beta, \text{trunc}) = 0$ if $M > \text{trunc}$ .

### Details

The beta-binomial distribution on model size is obtained by assigning each variable inclusion indicator independent Bernoulli distributions with probability  $w$ , and then giving  $w$  a  $\text{beta}(\alpha, \beta)$  distribution. Marginalizing over  $w$  leads to the number of included predictors having a beta-binomial distribution. The default hyperparameters lead to a uniform distribution over model size. The Truncated version assigns zero probability to all models of size  $> \text{trunc}$ .

### Value

returns an object of class "prior", with the family and hyperparameters.

### Author(s)

Merlise Clyde

### See Also

[bas.lm](#), [Bernoulli](#), [uniform](#)

Other priors modelpriors: [Bernoulli.heredity\(\)](#), [Bernoulli\(\)](#), [beta.binomial\(\)](#), [tr.poisson\(\)](#), [tr.power.prior\(\)](#), [uniform\(\)](#)



**Examples**

```

tr.beta.binomial(1, 10, 5)
library(MASS)
data(UScrime)
UScrime[, -2] <- log(UScrime[, -2])
crime.bic <- bas.lm(y ~ .,
  data = UScrime, n.models = 2^15, prior = "BIC",
  modelprior = tr.beta.binomial(1, 1, 8),
  initprobs = "eplogp"
)

```

tr.poisson

*Truncated Poisson Prior Distribution for Models***Description**

Creates an object representing the prior distribution on models for BAS using a truncated Poisson Distribution on the Model Size

**Usage**

```
tr.poisson(lambda, trunc)
```

**Arguments**

lambda	parameter in the Poisson distribution representing expected model size with infinite predictors
trunc	parameter that determines truncation in the distribution i.e. $P(M; \lambda, \text{trunc}) = 0$ if $M > \text{trunc}$

**Details**

The Poisson prior distribution on model size is obtained by assigning each variable inclusion indicator independent Bernoulli distributions with probability  $w$ , and then taking a limit as  $p$  goes to infinity and  $w$  goes to zero, such that  $p*w$  converges to  $\lambda$ . The Truncated version assigns zero probability to all models of size  $M > \text{trunc}$ .

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[bas.lm](#), [Bernoulli](#), [uniform](#)

Other priors modelpriors: [Bernoulli.hereditiy\(\)](#), [Bernoulli\(\)](#), [beta.binomial\(\)](#), [tr.beta.binomial\(\)](#), [tr.power.prior\(\)](#), [uniform\(\)](#)

**Examples**

```
tr.poisson(10, 50)
```

---

<code>tr.power.prior</code>	<i>Truncated Power Prior Distribution for Models</i>
-----------------------------	--

---

**Description**

Creates an object representing the prior distribution on models for BAS using a truncated Distribution on the Model Size where the probability of  $\gamma = p^{\text{kappa}} \text{lgammal}$  where  $\gamma$  is the vector of model indicators

**Usage**

```
tr.power.prior(kappa = 2, trunc)
```

**Arguments**

<code>kappa</code>	parameter in the prior distribution that controls sparsity
<code>trunc</code>	parameter that determines truncation in the distribution i.e. $P(\gamma; \alpha, \beta, \text{trunc}) = 0$ if $\text{lgammal} > \text{trunc}$ .

**Details**

The beta-binomial distribution on model size is obtained by assigning each variable inclusion indicator independent Bernoulli distributions with probability  $w$ , and then giving  $w$  a  $\text{beta}(\alpha, \beta)$  distribution. Marginalizing over  $w$  leads to the number of included predictors having a beta-binomial distribution. The default hyperparameters lead to a uniform distribution over model size. The Truncated version assigns zero probability to all models of size  $> \text{trunc}$ .

**Value**

returns an object of class "prior", with the family and hyperparameters.

**Author(s)**

Merlise Clyde

**See Also**

[bas.lm](#), [Bernoulli](#), [uniform](#)

Other priors modelpriors: [Bernoulli.heredity\(\)](#), [Bernoulli\(\)](#), [beta.binomial\(\)](#), [tr.beta.binomial\(\)](#), [tr.poisson\(\)](#), [uniform\(\)](#)

**Examples**

```
tr.power.prior(2, 8)
library(MASS)
data(UScrime)
UScrime[, -2] <- log(UScrime[, -2])
crime.bic <- bas.lm(y ~ .,
  data = UScrime, n.models = 2^15, prior = "BIC",
  modelprior = tr.power.prior(2, 8),
  initprobs = "eplogp"
)
```

trCCH

*Truncated Compound Confluent Hypergeometric function***Description**

Compute the Truncated Confluent Hypergeometric function from Li and Clyde (2018) which is the normalizing constant in the tcch density of Gordy (1998) with integral representation:

**Usage**

```
trCCH(a, b, r, s, v, k, log = FALSE)
```

**Arguments**

a	a > 0
b	b > 0
r	r >= 0
s	arbitrary
v	0 < v
k	arbitrary
log	logical indicating whether to return values on the log scale; useful for Bayes Factor calculations

**Details**

$$\text{tr.cch}(a,b,r,s,v,k) = \int_0^1 \frac{1}{v} u^{a-1} (1-vu)^{b-1} (k + (1-k)vu)^{-r} \exp(-s u) du$$

This uses a more stable method for calculating the normalizing constant using R's 'integrate' function rather than the version in Gordy 1998. For calculating Bayes factors that use the 'trCCH' function we recommend using the 'log=TRUE' option to compute log Bayes factors.

**Author(s)**

Merlise Clyde (<clyde@duke.edu>)

**References**

Gordy 1998 Li & Clyde 2018

**See Also**

Other special functions: [hypergeometric1F1\(\)](#), [hypergeometric2F1\(\)](#), [phi1\(\)](#)

**Examples**

```
# special cases
# trCCH(a, b, r, s=0, v = 1, k) is the same as
# 2F1(a, r, a + b, 1 - 1/k)*beta(a, b)/k^r

k = 10; a = 1.5; b = 2; r = 2;
trCCH(a, b, r, s=0, v = 1, k=k) *k^r/beta(a,b)
hypergeometric2F1(a, r, a + b, 1 - 1/k, log = FALSE)

# trCCH(a,b,0,s,1,1) is the same as
# beta(a, b) 1F1(a, a + b, -s, log=FALSE)
s = 3; r = 0; v = 1; k = 1
beta(a, b)*hypergeometric1F1(a, a+b, -s, log = FALSE)
trCCH(a, b, r, s, v, k)

# Equivalence with the Phi1 function
a = 1.5; b = 3; k = 1.25; s = 400; r = 2; v = 1;

phi1(a, r, a + b, -s, 1 - 1/k, log=FALSE)*(k^-r)*gamma(a)*gamma(b)/gamma(a+b)
trCCH(a,b,r,s,v,k)
```

---

uniform

*Uniform Prior Distribution for Models*

---

**Description**

Creates an object representing the prior distribution on models for BAS.

**Usage**

```
uniform()
```

**Details**

The Uniform prior distribution is a commonly used prior in BMA, and is a special case of the independent Bernoulli prior with  $\text{probs}=.5$ . The implied prior distribution on model size is  $\text{binomial}(p, .5)$ .

**Value**

returns an object of class "prior", with the family name Uniform.

**Author(s)**

Merlise Clyde

**See Also**

[bas.lm](#), [beta.binomial](#), [Bernoulli](#),

Other priors modelpriors: [Bernoulli.heredity\(\)](#), [Bernoulli\(\)](#), [beta.binomial\(\)](#), [tr.beta.binomial\(\)](#), [tr.poisson\(\)](#), [tr.power.prior\(\)](#)

**Examples**

```
uniform()
```

---

```
update.bas
```

*Update BAS object using a new prior*

---

**Description**

Update a BMA object using a new prior distribution on the coefficients.

**Usage**

```
## S3 method for class 'bas'
update(object, newprior, alpha = NULL, ...)
```

**Arguments**

object	BMA object to update
newprior	Update posterior model probabilities, probne0, shrinkage, logmarg, etc, using prior based on newprior. See <a href="#">bas</a> for available methods
alpha	optional new value of hyperparameter in prior for method
...	optional arguments

**Details**

Recomputes the marginal likelihoods for the new methods for models already sampled in current object.

**Value**

A new object of class BMA

**Author(s)**

Merlise Clyde <clyde@stat.duke.edu>

**References**

Clyde, M. Ghosh, J. and Littman, M. (2010) Bayesian Adaptive Sampling for Variable Selection and Model Averaging. *Journal of Computational Graphics and Statistics*. 20:80-101  
doi: [10.1198/jcgs.2010.09049](https://doi.org/10.1198/jcgs.2010.09049)

**See Also**

[bas](#) for available methods and choices of alpha

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [variable.names.pred.bas\(\)](#)

**Examples**

```
## Not run:
library(MASS)
data(UScrime)
UScrime[, -2] <- log(UScrime[, -2])
crime.bic <- bas.lm(y ~ ., data=UScrime, n.models=2^15, prior="BIC", initprobs= "eplogp")
crime.ebg <- update(crime.bic, newprior="EB-global")
crime.zs <- update(crime.bic, newprior="ZS-null")

## End(Not run)
```

---

```
variable.names.pred.bas
```

*Extract the variable names for a model from a BAS prediction object*

---

**Description**

S3 method for class 'pred.bas'. Simple utility function to extract the variable names. Used to print names for the selected models using estimators for 'HPM', 'MPM' or 'BPM'. for the selected model created by predict for BAS objects.

**Usage**

```
## S3 method for class 'pred.bas'
variable.names(object, ...)
```

**Arguments**

object	a BAS object created by predict from a BAS 'bas.lm' or 'bas.glm' object
...	other arguments to pass on

**Value**

a character vector with the names of the variables included in the selected model; in the case of 'BMA' this will be all variables

**See Also**

[predict.bas](#)

Other predict methods: [fitted.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#)

Other bas methods: [BAS](#), [bas.lm\(\)](#), [coef.bas\(\)](#), [confint.coef.bas\(\)](#), [confint.pred.bas\(\)](#), [diagnostics\(\)](#), [fitted.bas\(\)](#), [force.heredity.bas\(\)](#), [image.bas\(\)](#), [plot.confint.bas\(\)](#), [predict.basglm\(\)](#), [predict.bas\(\)](#), [summary.bas\(\)](#), [update.bas\(\)](#)

**Examples**

```
data(Hald)
hald.gprior = bas.lm(Y~ ., data=Hald, prior="ZS-null", modelprior=uniform())
hald.bpm = predict(hald.gprior, newdata=Hald[1,],
                  se.fit=TRUE,
                  estimator="BPM")
variable.names(hald.bpm)
```

---

which.matrix

*Coerce a BAS list object of models into a matrix.*

---

**Description**

This function coerces the list object of models to a matrix and fill in the zeros to facilitate other computations.

**Usage**

```
which.matrix(which, n.vars)
```

**Arguments**

which            a 'bas' model object x\$which  
n.vars            the total number of predictors, x\$n.vars

**Details**

which.matrix coerces x\$which into a matrix.

**Value**

a matrix representation of x\$which, with number of rows equal to the length of which.models or total number of models and number of columns x\$n.vars

**Author(s)**

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**See Also**

[bas](#)

Other `as.matrix` methods: [list2matrix.bas\(\)](#), [list2matrix.which\(\)](#)

**Examples**

```
data(Hald)
Hald.bic <- bas.lm(Y ~ ., data=Hald, prior="BIC", initprobs="eplogp")
# matrix of model indicators
models <- which.matrix(Hald.bic$which, Hald.bic$n.vars)
```



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